

# Comparison of frequency domain measures based on spectral decomposition for spontaneous baroreflex sensitivity assessment after Acute Myocardial Infarction

5 *Riccardo Pernice<sup>a,\*§</sup>, Laura Sparacino<sup>a,§</sup>, Giandomenico Nollo<sup>b</sup>, Alessandro Busacca<sup>a</sup>,  
Luca Faes<sup>a</sup>*

*<sup>a</sup>Department of Engineering, University of Palermo, 90128 Palermo, Italy*  
10 *<sup>b</sup>Department of Industrial Engineering, University of Trento, 38123 Trento, Italy*  
*\*[riccardo.pernice@unipa.it](mailto:riccardo.pernice@unipa.it)*

*§These authors contributed equally to this work*

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## Abstract

15 The **objective** of this study is to test the hypothesis that the baroreflex control of  
arterial pressure is altered after acute myocardial infarction (AMI) comparing  
different approaches for its evaluation based on the spectral analysis of spontaneous  
systolic arterial pressure (SAP) and heart period (HP) variability. We present a new  
20 **method** to assess directed interactions between SAP and RR from their linear model  
representation, based on pole decomposition of the model transfer function and  
evaluation of causal measures of coupling and gain from the poles associated to low  
frequency (0.04-0.15 Hz) oscillatory components. The method was compared with  
traditional non-causal approaches for the spectral analysis of the baroreflex gain, and  
with causal approaches based on the directed coherence, in a group of AMI patients  
25 and in Young and Old healthy controls studied at rest and during head-up tilt.  
Analysis of feedforward interactions from RR to SAP was also performed. Our  
**results** support the importance of using causal approaches to quantify separately  
baroreflex and feedforward interactions between RR and SAP. In AMI patients, these  
approaches revealed lower coupling and gain from SAP to RR, suggesting weaker  
30 effectiveness and lower sensitivity of the baroreflex after infarction, while they did  
not indicate clear alterations in the response to tilt. The postural stress altered instead  
the feedforward interactions selectively across groups, being related to decreased  
coupling only in Young and to increased gain mostly in AMI. These results have  
**significance** for the clinical assessment of the baroreflex and the physiological  
35 evaluation of cardiovascular interactions.

*Keywords:* spectral decomposition, frequency domain, causality, baroreflex, head-up  
tilt, cardiovascular control, acute myocardial infarction

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## 1. Introduction

The baroreflex mechanism has a key role in the short-term regulation of systolic arterial pressure (SAP) and heart period (measured as the RR interval of the electrocardiogram, ECG), which are known to dynamically interact in a closed loop, as a consequence of the baroreflex feedback of SAP on RR and of feedforward pathways of mechanical nature, whereby SAP is influenced by previous RR changes [1,2]. The baroreflex represents a fundamental mechanism to maintain the optimal blood pressure level continuously or in response to changing physiological conditions, and accomplishes to such a task modulating the heart rate [3–6]. Specifically, a decrease in arterial blood pressure evokes a baroreflex response leading to an increased heart rate, while an increase in arterial blood pressure is followed by the opposite effect. While in healthy subjects the baroreflex acts in response to physiological stressors such as change of posture, mental workload and others, an impairment of baroreflex control is thought to be often associated with age, orthostatic intolerance and pathologies like heart failure or acute myocardial infarction (AMI), as postural circulatory stress and cardiovascular diseases elicit baroreceptor unloading [7–10]. Therefore, assessment of the baroreflex sensitivity (BRS), often evaluated from the spontaneous beat-to-beat fluctuations of RR and SAP as the magnitude of the reflex change in RR corresponding to a unitary change in SAP, can provide valuable information for the analysis of cardiovascular regulation in normal and pathological conditions and can have an important diagnostic and clinical value [4,11,12]. An appropriate evaluation of the BRS should take into account not only the oscillatory nature of cardiovascular parameters, being able to separate contributions occurring in different frequency bands - typically divided into very low-frequency (VLF, up to 0.04 Hz), low-frequency (LF, 0.04 – 0.15 Hz) and high-frequency (HF, 0.15– 0.4 Hz) bands- but also the closed-loop nature of cardiovascular interactions [13]. Usually, the frequency analysis of feedback baroreflex interactions is carried out focusing on the LF band to avoid the confounding effects of other variables operating at different frequencies (e.g., respiration) and causal analysis methods are adopted to minimize the effect of non-baroreflex mechanisms on the BRS estimates [14–18]. On the other hand, though much less investigated, the feedforward mechanism from RR to SAP is also important in the assessment of the balanced cardiovascular regulation in normal and diseased conditions [7].

The directed (causal) coherence (DC) has been proposed and used as a linear frequency domain measure of causal interactions between coupled dynamic processes [11,19]. The DC from a source to a target process is computed from the spectral representation of their descriptive vector autoregressive (AR) model, whose parameters provide the basis to separate the power spectral density (PSD) of the target process into partial spectra related to its own dynamics and to the dynamics of the source process. As in practical applications both the DC function and baroreflex gain need to be quantified in specific frequency regions, empirical approaches are generally adopted. They consist in taking the maximum or average value, within the band of interest, of the analyzed spectral function. However, these approaches often lead to ambiguous choices (a maximum value can be absent within the observed band) or imprecise quantifications (the average value may be affected by spectral effects of nearby broadband oscillations) [11].

90 To overcome these limitations, the present study introduces a modification of the  
 causal coherence and of the previous definitions of spectral baroreflex gain [10,20]  
 based on the spectral decomposition method [21]. Specifically, spectral  
 decomposition is applied to the partial spectra of the PSD of the target process,  
 representing each partial spectrum as the sum of bell-shaped functions with features  
 (power, frequency, spectral bandwidth) related to the type and location (modulus and  
 95 phase) of the poles of the transfer function which defines the vector AR process in the  
 Z-domain. Then, the power content associated to the decomposed spectral  
 components with frequency inside the band of interest (here, the LF band) is  
 elaborated to obtain pole-specific measures of coupling and gain within the band.  
 These measures, which we refer to as “local” due to their frequency-specific nature,  
 are compared in the present work with the corresponding “global” measures obtained  
 as the band-averaged DC and gains. The analysis, also extended to the comparison  
 100 between causal and non-causal measures of gain and to the quantification of  
 feedforward (non-baroreflex) interactions, is performed on the joint variability series  
 of the heart period and the systolic arterial pressure measured in a group of post-AMI  
 patients monitored at rest and during orthostatic stress, as well as in two control  
 groups of healthy subjects (younger and age-matched with AMI). Preliminary  
 105 methodological and applicative results have been presented in a reduced form in  
 conference contributions [13], [22].

## 2. Methods

### 2.1 Measures of causal coupling and gain derived from parametric cross-spectral analysis

110 Let us consider a bivariate stochastic process  $Y$  composed by two jointly stationary,  
 zero mean discrete stochastic processes  $y_1$  and  $y_2$ . Defining  $Y(n) = [y_1(n) \ y_2(n)]^T$   
 as the vector variable sampling the process at time  $t_n = nT_s$ , where  $T_s$  is the sampling  
 period, it is possible to express the causal interactions occurring between the  
 processes in a parametric form through a  $p$ -order bivariate autoregressive (AR) model  
 115 as follows [22–24]

$$Y(n) = \sum_{k=0}^p \mathbf{A}(k)Y(n-k) + U(n), \quad (1)$$

120 being  $U(n) = [u_1(n) \ u_2(n)]^T$  a vector of zero-mean uncorrelated white noises with  
 $2 \times 2$  diagonal covariance matrix  $\mathbf{\Sigma} = \text{diag}\{\sigma_1^2, \sigma_2^2\}$ , and  $\mathbf{A}(k)$  the  $2 \times 2$  coefficient  
 matrix in which the coefficient  $a_{ij}(k)$  describes the interaction from  $y_j(n-k)$  to  
 $y_i(n)$  ( $i, j=1, 2$ ).

The estimated model coefficients are represented in the Z domain through the Z-  
 transform of (1), thus yielding  $Y(z) = \mathbf{H}(z)U(z)$ , where the  $2 \times 2$  transfer matrix is

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$$\mathbf{H}(z) = \begin{bmatrix} H_{11}(z) & H_{12}(z) \\ H_{21}(z) & H_{22}(z) \end{bmatrix} = [\mathbf{I} - \mathbf{A}(z)]^{-1} = \bar{\mathbf{A}}(z)^{-1}, \quad (2)$$

130 being  $\mathbf{A}(z) = \sum_{k=1}^p \mathbf{A}(k)z^{-k}$  the coefficient matrix in the  $Z$  domain and  $\mathbf{I}$  the  $2 \times 2$  identity matrix. Taking the inverse of  $\bar{\mathbf{A}}(z)$ , each element of the transfer matrix is represented as follows ( $i, j=1, 2; i \neq j$ )

$$H_{ii}(z) = \frac{\bar{A}_{jj}(z)}{|\bar{\mathbf{A}}(z)|}, \quad H_{ij}(z) = \frac{-\bar{A}_{ij}(z)}{|\bar{\mathbf{A}}(z)|}. \quad (3)$$

135 Computing  $\mathbf{H}(z)$  on the unit circle in the complex plane ( $\mathbf{H}(f) = \mathbf{H}(z)|_{z=e^{j2\pi fT_s}}$ ), the  $2 \times 2$  spectral density matrix of the bivariate process in the frequency domain becomes  $\mathbf{S}(f) = \mathbf{H}(f)\mathbf{\Sigma}\mathbf{H}^*(f)$ , where  $*$  indicates the Hermitian transpose. In this matrix, the diagonal terms  $S_{ii}(f)$  correspond to the auto-spectra, while the off-diagonal terms  $S_{ij}(f)$  correspond to the cross-spectra.

140 From the frequency domain representation of the AR model, the diagonal elements of the spectral density matrix can be elaborated to estimate a non-causal measure of spectral gain from  $y_j$  to  $y_i$  ( $i, j = 1, 2, i \neq j$ ):

$$G_{ij}^\alpha(f) \triangleq \sqrt{\frac{S_{ii}(f)}{S_{jj}(f)}} \quad (4)$$

145 where the superscript  $\alpha$  denotes the non-causal index first used by Pagani et al. [20]. The  $\alpha$  measure is a non-causal index of gain, because its main assumption is that the whole variability of  $y_i$  is driven by  $y_j$ . To get a causal measure, first each auto-spectrum is expressed as the sum of causal contributions from Eqs. (2,3) to yield

$$150 \quad S_i(f) \triangleq S_{ii}(f) = \sigma_i^2 |H_{ii}(f)|^2 + \sigma_j^2 |H_{ij}(f)|^2, \quad (5)$$

being  $\sigma_j^2 |H_{ij}(f)|^2 \triangleq S_{i|j}(f)$  the partial spectrum of  $y_i$  given  $y_j$  ( $i, j=1, 2$ ). A left-side normalization of (5) produces  $\gamma_{ii}^2(f) + \gamma_{ij}^2(f) = 1$ , where

$$155 \quad \gamma_{ij}^2(f) \triangleq \frac{\sigma_j^2 |H_{ij}(f)|^2}{S_{ii}(f)} = \frac{S_{i|j}(f)}{S_i(f)} \quad (6)$$

160 is the squared directed (causal) coherence (DC) from  $y_j$  to  $y_i$ , a function assessing the normalized coupling strength from  $y_j$  to  $y_i$  in the frequency domain [19]. The DC ranges between 0 and 1, being 0 when  $y_j$  does not cause  $y_i$  at frequency  $f$ , and 1 when the whole power of  $y_i$  at frequency  $f$  is due to the variability of  $y_j$  [24]. The causal information conveyed in the DC allows to define a causal measure of spectral gain, first used by Faes et al. [10]:

$$165 \quad G_{ij}^\gamma(f) \triangleq G_{ij}^\alpha(f)\gamma_{ij}^2(f) \quad (7)$$

The  $\gamma$  measure (7) weights the power ratio defined in (4) through the causal coherence from  $y_j$  to  $y_i$ .

## 2.2 Measures of causal coupling and gain derived from causal spectral decomposition

170 The measures defined by Eqs. (4, 6, 7) are “frequency-specific” in the sense that they are computed at each frequency; therefore, when they have to be computed within a specific frequency band, an “overall” estimate is typically extracted as the average of that measure within the band. For this reason, these are referred to as “global measures” in the following. Nonetheless, such global values in the selected  
175 band can be misleading, being unable to objectively quantify the causal contribution of the source process to the power of the target one [22]. As an alternative, we herein propose “pole-specific” measures that can be related to the poles of the transfer function of the bivariate AR model. Their computation is associated to the pole frequency, and thus allows to get “local” measures of causal coupling or gain, where  
180 “local” is meant to indicate individual oscillations localized at specific frequencies. In the following, we define local measures of causal coupling (pole-specific spectral causality, PSSC [22]) and causal and non-causal local measures of gain (pole-specific spectral gain, PSSG) to assess in the frequency domain both the strength and the magnitude of the directed interactions between two processes.

185 Exploiting spectral decomposition [21], each transfer function defined as in (3) is decomposed as the sum of  $q$  spectral components ( $q \cong p/2$ ), which correspond to the poles of the determinant of  $\bar{\mathbf{A}}(z)$ . Every spectral component is described by a specific profile and has an associated frequency (related to the pole phase) and power (related to the pole residual). In this way, the complex partial PSD of the  $i^{th}$  process given the  
190  $j^{th}$  process, which can be written in the Z-domain as

$$S_{i|j}(z) = H_{ij}(z)\sigma_j^2 H_{ij}^*(1/z^*), \quad (8)$$

can be expanded decomposing the  $i - j^{th}$  transfer function as

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$$H_{ij}(z) = \frac{\bar{A}_{ij}(z)}{|\bar{\mathbf{A}}(z)|} = \frac{\bar{A}_{ij}(z)}{\prod_k(z-z_k)}, \quad (9)$$

being the poles  $z_k$ ,  $k=1, \dots, q$  the roots of  $|\bar{\mathbf{A}}(z)|$ . The expansion of each partial spectrum in (8) is performed exploiting Heaviside decomposition with simple  
200 fractions relevant to all its poles (i.e., the poles  $z_k$  inside the unit circle and their reciprocals  $\bar{z}_k = z_k^{-1}$  outside the unit circle, with  $k=1, \dots, q$ ), which are fractions weighted by the relevant residuals of  $S_{i|j}(z)$  (i.e.,  $r_k z_k$  and  $-r_k \bar{z}_k$ ), to get [21]

$$S_{i|j}(z) = \sum_{k=1}^q S_{i|j}^{(k)}(z), \quad S_{i|j}^{(k)}(z) = \frac{r_k z_k}{z-z_k} - \frac{r_k \bar{z}_k}{z-\bar{z}_k}. \quad (10)$$

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After extracting the residuals and expanding the partial spectrum in simple fractions and given that  $S_i(z) = S_{i|i}(z) + S_{i|j}(z)$ , we obtain the spectrum of  $y_i$  computing (10) for values of  $z$  on the unit circle of the complex plane [22]:

$$S_i(f) = \sum_{k=1}^q S_i^{(k)}(f) = \sum_{k=1}^q S_{i|i}^{(k)}(f) + S_{i|j}^{(k)}(f). \quad (11)$$

The  $k^{th}$  spectral component  $S_{i|j}^{(k)}(f)$ ,  $i, j = 1, 2$ , has an associated frequency related to the pole phase,  $f(k) = \arg\{z_k\}$ , and power related to the pole residuals,  $P_{i|j}(k) = r_k$  for real poles and  $P_{i|j}(k) = r_k + r_k^*$  for complex conjugate poles. It is then possible to achieve a decomposition for the DC from  $y_j$  to  $y_i$  normalizing the spectral components to the total spectrum as follows:

$$\gamma_{ij}^2(f) = \sum_{k=0}^q \gamma_{ij}^{2(k)}(f), \quad \gamma_{ij}^{2(k)}(f) \triangleq \frac{S_{i|j}^{(k)}(f)}{S_i(f)}. \quad (12)$$

Furthermore, causal contributions to the spectral power can be obtained by integrating each spectral component over the whole frequency axis. This allows to exploit (11) for decomposing the variance of the process  $y_i$ ,  $\lambda_i^2$ , as

$$\lambda_i^2 = \frac{2}{f_c} \int_0^{f_c} S_i(f) df = \sum_{k=0}^q P_{i|i}(k) + P_{i|j}(k) = \sum_{k=0}^q P_i(k) \quad (13)$$

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where  $P_{i|i}(k) = \int_0^{f_c} S_{i|i}^{(k)}(f) df$  is the part of the variance of  $y_i$  due to its own dynamics and relevant to  $k^{th}$  oscillation (pole),  $P_{i|j}(k) = \int_0^{f_c} S_{i|j}^{(k)}(f) df$  is the part of the variance of  $y_i$  due to  $y_j$  and relevant to the  $k^{th}$  pole, and by summing these two contributions to the variance we get the part of the variance of  $y_i$  relevant to the  $k$ -th pole, *i.e.*,  $P_i(k) = P_{i|i}(k) + P_{i|j}(k)$ . Using these partial variances, the PSSC relevant to the  $k^{th}$  oscillation is obtained as a local causal measure of coupling from  $y_j$  to  $y_i$ :

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$$\gamma_{i|j}^2(k) \triangleq \frac{P_{i|j}(k)}{P_i(k)}. \quad (14)$$

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The PSSC ranges between 0 and 1, being equal to 0 when the power of the  $k^{th}$  oscillation of  $y_i$  (*i.e.* the oscillation at frequency  $f_k$ ) is totally due to its internal dynamics and equal to 1 when it is totally caused by the dynamics of  $y_j$  assessed at the same frequency  $f_k$ . Given that the frequency  $f_k$  is associated to a specific causal spectral profile, the corresponding PSSC value represents an objective measure of the causal power at that frequency, so that the total causal power in a specific frequency band  $f$  can be easily computed summing all PSSC values with frequency within that band. Using the same formalism, we also define a local non-causal measure of gain (local PSSG) from  $y_j$  to  $y_i$  related to the  $k^{th}$  pole as

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$$G_{ij}^\alpha(k) \triangleq \sqrt{\frac{P_i(k)}{P_j(k)}}. \quad (15)$$

250 The gain measure defined in (15) relates the whole variability of the output process  $y_i$  to that of the input process  $y_j$  without attempting to separate causal and non-causal contributions. To get a causal measure, we consider only the power of  $y_i$  causally due to  $y_j$  and related to the  $k^{th}$  pole and define the local PSSG from  $y_j$  to  $y_i$  as:

$$G_{i|j}^{\gamma}(k) = \sqrt{\frac{P_{i|j}(k)}{P_j(k)}}. \quad (16)$$

255 Note that, while the spectral causality is an adimensional measure, the spectral gain is expressed in units of measurement of the output series divided by units of measurement of the input series.

### 2.3 Computation in cardiovascular variability analysis

260 To evaluate the new proposed local measures in comparison with the traditional global measures of both causal coupling and gain, the latter distinguishing also between causal and non-causal indices, we considered pairs of simultaneously observed beat-to-beat time series of SAP and RR, corresponding respectively to realizations of the processes  $y_1$  and  $y_2$ . In the frequency domain analysis, the coupling and gain functions were evaluated within the LF band, ranging from  $f_{inf}^{LF} = 0.04$  Hz to  $f_{sup}^{LF} = 0.15$  Hz [25], in order to minimize the effects of non-baroreflex mechanisms on the assessed measures and especially to avoid the confounding effects of respiration on SAP and RR which are primarily confined in the HF band [14–18,25]. Accordingly, after computing the spectrum of RR and SAP as well as their partial spectra and decomposition, the global measures of causal coupling (6), non-causal gain (4) and causal gain (7) were averaged in the LF band to get the following indexes ( $i, j = 1, 2, i \neq j$ ):

- Global causal coupling:  $\gamma_{ji}^2(LF) \triangleq \frac{1}{f_{sup}^{LF} - f_{inf}^{LF}} \int_{f_{inf}^{LF}}^{f_{sup}^{LF}} \gamma_{ji}^2(f) df$  (17)

- Global non-causal gain:  $G_{ji}^{\alpha}(LF) \triangleq \frac{1}{f_{sup}^{LF} - f_{inf}^{LF}} \int_{f_{inf}^{LF}}^{f_{sup}^{LF}} G_{ji}^{\alpha}(f) df$  (18)

- Global causal gain:  $G_{ji}^{\gamma}(LF) \triangleq \frac{1}{f_{sup}^{LF} - f_{inf}^{LF}} \int_{f_{inf}^{LF}}^{f_{sup}^{LF}} G_{ji}^{\gamma}(f) df$ . (19)

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On the other hand, local band-specific measures were obtained applying the equations for causal coupling (14), non-causal gain (15) and causal gain (16) after summing the power content of all the components with central frequency  $f_k$  within the LF range, *i.e.* computing ( $i, j = 1, 2, i \neq j$ ):

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- Local causal coupling:  $\gamma_{j|i}^2(k_{LF}) = \frac{\sum_{f_k \in LF} P_{j|i}(k)}{\sum_{f_k \in LF} P_j(k)}$  (20)

- Local non-causal gain:  $G_{ji}^\alpha(k_{LF}) = \sqrt{\frac{\sum_{f_k \in LF} P_{j|k}(k)}{\sum_{f_k \in LF} P_{i|k}(k)}}$  (21)

- Local causal gain:  $G_{ji}^\gamma(k_{LF}) = \sqrt{\frac{\sum_{f_k \in LF} P_{j|i}(k)}{\sum_{f_k \in LF} P_{i|k}(k)}}$ . (22)

285 The derivation of the local measures of coupling and gain is illustrated in Fig. 1 for representative SAP and RR time series. For the same time series, the derivation of the local measures of coupling and gain is illustrated in Fig. 2.

#### 2.4 Experimental protocol and data analysis

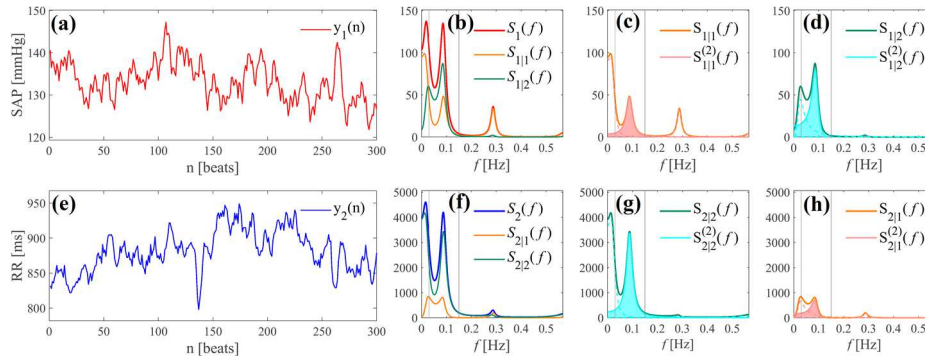
290 The study included 35 post-AMI patients ( $58.5 \pm 10.2$  yrs, 4 female), examined  $10 \pm 3$  days after AMI, and two groups of healthy subjects, 19 young ( $25.0 \pm 2.6$  yrs, 9 female) and 12 old ( $63.1 \pm 8.3$  yrs, 9 female), all monitored in the resting supine position and in the  $60^\circ$  upright position reached after passive head-up tilt [7,26]. After recording ECG (Siemens Mingograph, hardware bandpass filter 0.3-1000Hz, lead II) and non-invasive finger arterial pressure (Ohmeda 2300; Englewood, CO), the beat-to-beat variability series of RR interval and SAP were measured on a beat-by-beat basis. Two stationary and artifact-free windows, each of  $\sim 5$  min duration, corresponding to 300 beats, were then selected in correspondence with the two epochs of the test. All signals were filtered to avoid the effect of long-term trends on the data analysis, employing an autoregressive high-pass filter with zero phase. Further details on experimental protocol and data acquisition can be found in [7,26].

300 Each pair of RR and SAP series were fitted by a bivariate AR model, after allowing instantaneous zero-lag effects in the direction from SAP to RR, (i.e., setting  $a_{21}(0) \neq 0$  and  $a_{ij}(0) = 0$  as a constraint for model identification) to allow fast within-beat baroreflex influences in agreement with the adopted measurement convention. Model identification was performed via the vector least-squares approach, setting the model order  $p$  according to the multivariate version of the Akaike criterion [27]. In some cases, the use of the Akaike criterion led to negative power contributions and/or complex coupling and gain indices as a result of the residue theorem, which may have caused erroneous final results. To avoid such a problem, all the obtained results were manually checked and, in case of negative power contributions, the model order  $p$  was slightly varied and thus manually selected to achieve positive power values. After AR identification, estimation of the global and local measures of causal coupling and causal and non-causal gain were computed from the estimated model parameters as described in the previous sections. Spectral analysis was performed assuming the series as uniformly sampled with the mean heart period  $\langle RR \rangle$  taken as the sampling period  $T_s$ , so that the Nyquist frequency in each spectral representation was taken as  $\frac{f_s}{2} = \frac{1}{2\langle RR \rangle}$ .

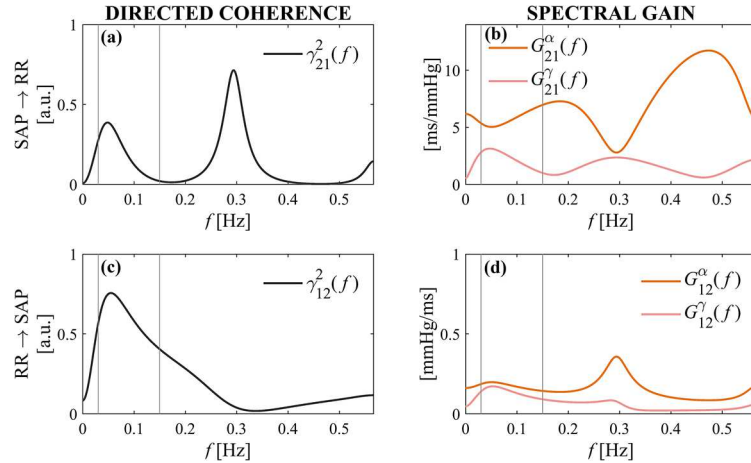
315 As regards the statistical analysis, the distributions of the coherence and gain indices were tested for normality using the Anderson-Darling test [28,29]. Since the



320 hypothesis of normality was rejected for most of the distributions, and given the small  
sample size especially for young and old subjects, non-parametric tests were  
employed [30]. For any given group, the statistical significance of the difference  
between rest and tilt conditions was assessed using the Wilcoxon signed rank test  
[31]. Afterwards, the statistical significance of the differences of the median of the  
325 distributions among groups at a given physiological condition (rest or tilt) was  
assessed using the non-parametric Kruskal Wallis test [32]. When the null hypothesis  
that the data in each group comes from the same distribution was rejected, a pairwise  
comparison was carried out using the Dunn post-hoc test with Šidák correction for  
multiple comparisons ( $m=3$ ) [33,34] to assess differences between group pairs (Young  
330 vs Old, Old vs AMI and Young vs AMI) at a given condition (rest or tilt). Finally, to  
assess the statistical differences between global and local or causal vs non-causal  
indices given the group and the condition, the Wilcoxon signed rank test [31] was  
employed. All statistical tests were carried out with 5% significance level.



335 Figure 1. Example of causal spectral decomposition of the interactions between systolic arterial pressure  
(SAP) and heart period (RR intervals). (a) and (e): SAP and RR time series, respectively, for a  
representative AMI subject in the resting supine position. (b): the power spectrum of SAP is given by  $S_1(f)$   
and is decomposed as the sum of a causal spectrum [ $S_{1|2}(f)$ ] and a non-causal spectrum [ $S_{1|1}(f)$ ]; (c) and  
340 (d): the non-causal (c) and causal (d) spectra of SAP are in turn decomposed with the spectral  
decomposition method into contributions associated to specific oscillations of the time series, with those in  
the LF band (pole  $k=2$ ) given by  $S_{1|1}^{(2)}(f)$  (non-causal part, c) and  $S_{1|2}^{(2)}(f)$  (causal part, d). (f): the power  
spectrum of RR is given by  $S_2(f)$  and is decomposed as the sum of a causal spectrum [ $S_{2|1}(f)$ ] and a non-  
causal spectrum [ $S_{2|2}(f)$ ]; (g) and (h): the non-causal (g) and causal (h) spectra of RR are in turn  
345 decomposed into contributions associated to specific oscillations, with those in the LF band given by  
 $S_{2|2}^{(2)}(f)$  (non-causal part, g) and  $S_{2|1}^{(2)}(f)$  (causal part, h). Local measures are computed as follows: the  
PSSC from SAP to RR in the LF band,  $\gamma_{21}^2(k_{LF})$ , is computed as the ratio between the power  $P_{2|1}^{(2)}(LF)$   
(pink area, (h)) and the total power  $P_{2|1}^{(2)}(LF) + P_{2|2}^{(2)}(LF)$  (pink + cyan areas, (h) and (g)); the non-causal  
PSSG from SAP to RR in the LF band,  $G_{21}^x(k_{LF})$ , is computed as the squared root of the ratio between the  
total power  $P_{2|1}^{(2)}(LF) + P_{2|2}^{(2)}(LF)$  (pink + cyan areas, (h) and (g)) and the total power  $P_{1|2}^{(2)}(LF) + P_{1|1}^{(2)}(LF)$   
350 (cyan + pink areas, (d) and (c)); the causal PSSG from SAP to RR in the LF band,  $G_{21}^y(k_{LF})$ , is computed as  
the squared root of the ratio between the power  $P_{2|1}^{(2)}(LF)$  (pink area, (h)) and the total power  $P_{1|2}^{(2)}(LF) + P_{1|1}^{(2)}(LF)$  (cyan + pink areas, (d) and (c)). The same procedure (LF) applies to the computation of the directed  
coherence and gain indexes from RR to SAP.



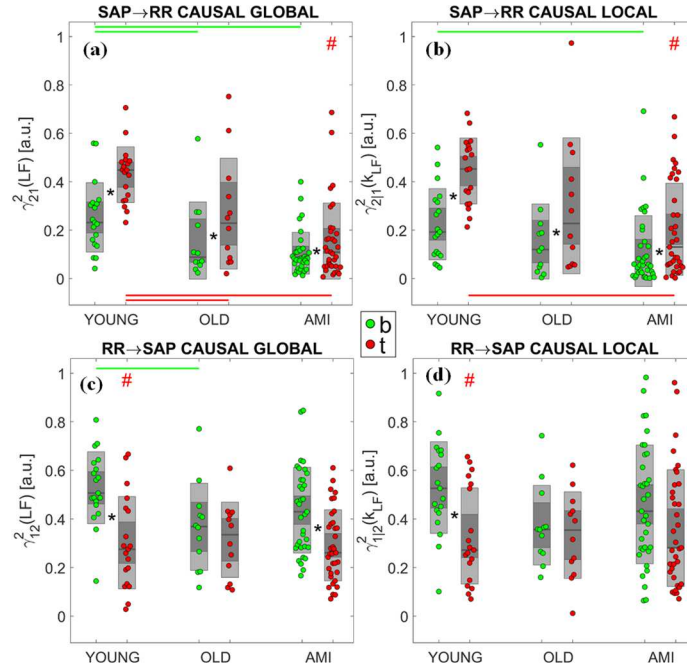
355 Figure 2. Example of frequency domain spectral analysis of baroreflex (SAP→RR, top row) and  
 feedforward (RR→SAP, bottom row) interactions for a representative AMI subject in the resting supine  
 position. (a) and (c): directed coherence function,  $\gamma_{ij}^2(f)$  ( $i,j=1,2$  and  $i \neq j$ ); (b) and (d): spectral profiles of  
 the non-causal gain  $G_{ij}^\alpha(f)$  and of the causal gain  $G_{ij}^\gamma(f)$  ( $i,j=1,2$  and  $i \neq j$ ). These ‘frequency-specific’  
 measures were averaged in the low frequency band (vertical grey lines in each plot) to get global measures.

### 360 3. Results

In this section, the results of LF spectral decomposition and analysis are reported,  
 showing the distributions of the causal and non-causal indices described in Section 2,  
 alongside with the outcomes of the statistical analyses carried out between different  
 physiological conditions and between the various measures. The results in terms of  
 365 causal coupling measures and of gain measures are depicted in Figures 3 and 4,  
 respectively.

In the baroreflex direction from SAP to RR (Fig. 3, top panels), both global and  
 local measures of causal coupling were significantly higher during tilt than during rest  
 for all groups, showing a clear response to postural stress of the DC along the  
 370 baroreflex. Moreover, while the global values were significantly lower in Old and  
 AMI subjects compared with Young during both rest and tilt (Fig. 3a), the local index  
 showed a significant decrease only in AMI compared with Young in both postural  
 conditions (Fig. 3b).

As regards the gain measures in the same direction (Fig. 4, top panels) the non-  
 375 causal indices decreased in Old compared to Young during postural stress, while they  
 did not elicit striking differences relevant to the AMI group (Fig. 4a,c). On the  
 contrary, the causal gain indices (measured both globally and locally) detected a  
 statistically significant decrease of the gain not only in Old, but also in AMI  
 compared to Young, in both body positions (Fig. 4b,d). Moreover, the comparison  
 380 between the two experimental conditions revealed that the gain index decreased  
 significantly moving from rest to tilt in all the three groups if computed using the non-  
 causal methods (Fig. 4a,c), only in AMI patients using the causal global method (Fig.  
 4b), and in none of the groups using the causal local method (Fig. 4d).



385 Figure 3. Results of low frequency (LF) spectral causality analysis of baroreflex (top row) and feedforward (bottom row) interactions based on the global and local directed (causal) coherence (respectively,  $\gamma_{ij}^2(LF)$  and  $\gamma_{ij}^2(k_{LF})$ , where  $ij=SAP, RR$  and  $i \neq j$ ). Plots depict the distributions across subjects, shown as individual values and box-plot distributions, of the directed coherence from SAP to RR and from RR to SAP, computed in the supine (green) and upright (red) body positions. Statistically significant differences: 390 \*, rest vs. tilt; #, global vs. local; -, YOUNG vs. OLD, YOUNG vs. AMI or OLD vs. AMI.

Global and local measures of coupling and gain were also investigated in the feedforward direction, where mechanical effects are known to alter RR variability according to changes in SAP variability [16]. The local measure of causal coupling 395 decreased significantly from rest to tilt in Young but not in Old and AMI (Fig. 3d). The corresponding global measure of feedforward coupling decreased with tilt also in the AMI patients, and was lower in Old than in Young in the supine position (Fig. 3c). As regards the gain from RR to SAP, the non-causal measures (both global and local) were significantly higher during orthostatic stress in all three groups (Fig. 4e,g), while the causal global measure increased significantly in Old and AMI (Fig. 4f) and the causal local measure increased significantly only in AMI (Fig. 4h).

Figs. 3 and 4 illustrate also that non-causal measures of gain are always significantly lower than the corresponding causal ones, which is an obvious consequence of their mathematical formulation (i.e., the global causal gain (7) is obtained multiplying the global non-causal gain (4) by the DC, and the local causal gain (16) contains at the numerator a fraction of the power contained in the global non-causal gain (15)). Moreover, local measures tend to be lower in value than the corresponding global ones, being related to a specific oscillation in the LF band; the statistically significant differences in Fig. 4 are just a few, i.e. for Young (tilt and rest) 405 and Old (tilt) in SAP  $\rightarrow$  RR non causal index, for Young (rest) with regard to 410

SAP→RR causal index and for Young and AMI (tilt) causal index in RR→SAP direction. Instead, comparing DC measures (Fig. 3), global and local indices resulted statistically different just for AMI tilt (SAP→RR direction) and for Young tilt (RR→SAP).

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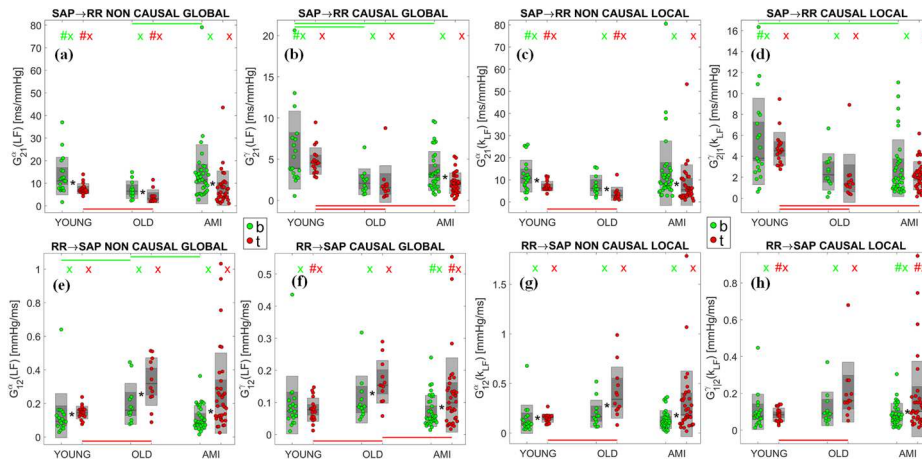


Figure 4. Distributions over subjects of the BRS computed in the LF band with the two different approaches (from left: non-causal and causal global measures, non-causal and causal local measures) along the two directions of interest (top: baroreflex direction from SAP to RR; bottom: feedforward direction from RR to SAP), in the rest (green) and tilt (red) phases of the testing protocol. Statistically significant differences: \*, rest vs. tilt; #, global vs. local; -, YOUNG vs. OLD, YOUNG vs. AMI or OLD vs. AMI; x, non-causal vs. causal.

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#### 4. Discussion

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Evaluation of the baroreflex gain is considered an important tool in clinical practice for diagnosis and prognosis in many cardiac diseases, including acute myocardial infarction [4–7,10,13]. A decreased baroreflex sensitivity has been already observed in several pathological conditions as a marker of cardiovascular system impairment [7,10,13]. In this study, we propose a new method to assess the BRS in the frequency domain, investigating the usefulness of using local causal measures (*i.e.* ‘pole-specific’ measures) in place of already widely employed global approaches (*i.e.* ‘frequency-specific’ measures); the comparison is extended to non-baroreflex (feedforward) interactions to investigate the relevant underlying mechanisms.

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Causal methods have been already proved in the literature as useful tools, typically more reliable than non-causal ones, for the quantitative assessment of cardiovascular regulatory mechanisms [10,35,36]. Moreover, preliminary results have shown that the frequency-averaging approaches commonly undertaken to obtain an individual value from a spectral function (e.g., the DC) for evaluating it within a specific band of interest (e.g., the LF band) may be inaccurate as they can incorporate spectral contributions originating from neighboring frequency ranges [22]. In the following,

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we compare more thoroughly causal vs non-causal and local vs global indices to put in evidence strengths and drawbacks of each approach in light of our results.

#### *4.1 Causal vs. non-causal assessment of cardiovascular interactions*

445 In physiological conditions, RR and SAP normally reciprocally affect each other due to both regulatory feedback and mechanical feedforward coupling mechanisms (mainly of mechanical nature, such as the Windkessel and Frank-Starling effects) [23,36]. The presence of a closed-loop interaction between the heart period and the systolic arterial pressure highlighted by past works [2,23,36,37] suggests the importance to implement causality to assess cardiovascular interactions.

450 Our results confirm the findings already reported in the literature [10,35,36] highlighting the suitability of a causal ( $\gamma$  index) instead of a non-causal method ( $\alpha$  index) for the frequency domain evaluation of BRS. The values of causal gain indices obtained in our analysis are significantly lower than the corresponding non-causal ones in all the groups of subjects during rest and head-up tilt. This finding confirms  
455 that the non-causal approach may overestimate the BRS, while closed-loop modeling allows to separately evaluate feedback and feedforward pathways, quantifying their relative contribution to the overall cardiovascular regulation [10].

The mixing between feedback and feedforward RR-SAP interactions can be also the reason why the non-causal measures of BRS detected a lower gain during postural stress not only in the AMI patients, but also in the Old and Young healthy subjects.  
460 The effect of tilt maneuver on the power spectrum of heart rate variability (HRV) is widely known [38–41], as is recognized that it evokes a greater effectiveness of the baroreflex that is mirrored by higher values of coupling between SAP and RR during postural stress; this effect was observed both in the present and in previous  
465 investigations [10,16,42]. Nevertheless, a decrease of the magnitude of the reflex, mirrored by the gain estimate, is more difficult to explain physiologically; here, the latter was observed only in the AMI patients using the global causal measure. Moreover, the causal measures of gain (both global and local) highlighted a significantly lower gain in the Old group, and especially in the AMI group, compared  
470 to the Young group. These decreased BRS values were observed in both the supine and upright position, documenting a reduced response of the baroreflex likely related to an impairment occurring with the coronary disease.

As regards the mechanical feedforward, the significant increase with head-up tilt observed for the non-causal measures of gain in all groups was found consistently for  
475 both causal measures only in the AMI patients. This result may be an indication of a physiological response to tilt that occurs as a consequence of the disease (as discussed in Sect 4.3), and is observed in young healthy subjects only when non-causal methods are improperly used to assess the gain function.

#### 4.2 *Global vs. local assessment of cardiovascular interactions*

480 From a methodological point of view, the operation of averaging spectral functions  
(like the DC or the gain) to get a global index within an assigned frequency band  
could lead to inaccurate evaluation of the index, because external broadband  
oscillations may convey information into the analyzed band. This aspect, that we  
demonstrated recently in a theoretical example [22], is a major deal in cardiovascular  
485 variability analysis where VLF oscillations are often predominant and may thus have  
an impact on the evaluation of the DC or the BRS in the LF band of the spectrum  
[22]; a similar effect of spreading between adjacent bands may involve also the HRV  
oscillations located in the HF band and mainly due to respiration [14,25,38]. On the  
other hand, the method of spectral decomposition allows to focus more objectively on  
490 the spectral content within specific ranges, lastly resulting in DC and gain values  
(local measures) which are expected to reflect more accurately the underlying  
mechanisms of effectiveness and magnitude of a reflex with regard to the oscillations  
of physiological interest.

In the light of the methodological considerations above, we interpret the agreement  
495 often found between global and local measures as indicative of a limited impact of  
broadband VLF or HF spectral contributions into the analyzed LF band. This was the  
case, in the feedback direction from SAP to RR, for the significantly lower values of  
both causal coupling and gain observed in AMI compared to Young in both body  
positions and for the increase with tilt of the causal coupling in all groups, and, in the  
500 feedforward direction from RR to SAP, for the decrease in Young (but not in Old and  
AMI) of the causal coupling and the increase in AMI (but not in Young) of the causal  
gain observed moving from rest to tilt. These results, consistently found using both  
global and local causal measures, are interpreted as robust and are discussed  
physiologically in Sect. 4.3.

505 On the other hand, a disagreement between global and local measures of causal  
influence is likely indicating an effect of bands external to the LF on the global  
computation based on averaging. In our analysis, the main occurrence of this  
disagreement is the detection in AMI of a significant decrease of the causal gain from  
SAP to RR observed moving from rest to tilt with the global method but not with the  
510 local method; this suggests that, in the analysed post-infarction patients, a depressed  
BRS response to tilt occurs with contributions to cardiovascular variability located in  
frequency bands other than the LF. The other differences observed between the global  
and local approaches regard mostly comparisons involving the Old group; these may  
be explained also by the small size of this group (see Sect. 4.4).

#### 515 4.3 *Characterization of feedback and feedforward cardiovascular interactions via a local causal approach*

As discussed in the previous subsections, the local causal method may be  
considered the methodologically most accurate approach for assessing the coupling  
and gain related to specific oscillations of the two analyzed time series, as it is able to  
520 separately evaluate feedback and feedforward pathways and to avoid confounding

spectral contributions from other frequency ranges. In the analyzed data, the local causal approach showed a tendency to detect less statistically significant differences in the comparisons between groups and conditions; as these seem to be more conceivable and with a more robust physiological meaning, we discuss them from a physiological viewpoint.

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Compared to the young healthy controls, a significantly lower causal coupling along the baroreflex direction from SAP to RR was detected in post-AMI patients during each of the two analyzed experimental conditions; the difference with Young was not statistically significant in the Old group when the local method was adopted to measure the causal coupling. This result is in agreement with previous findings in the literature which highlighted an overall lower synchronization index and an increase of the number of subjects showing uncoupled RR and SAP dynamics [7]. Moreover, the post-AMI patients showed also significantly an impaired baroreflex gain compared to the young subjects, both at rest and during tilt. This result, that was observed also in Old during tilt, can be attributed to the higher sympathetic tone of these subjects and to an inability to respond to changes in cardiac output by further sympathetic activation [7], and can be related to the known reduction of heart rate variability typically occurring in elderly and AMI subjects [43–45].

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When the response to head-up tilt was analyzed, we observed in all groups a significant increase of the causal coupling along the baroreflex pathway moving from the supine to the upright position. This finding is typically related to the sympathetic activation produced by the postural stress induced by tilt [46,47] which has already been widely observed in the literature [38,39,41,48]. Overall, it reflects the increased effectiveness of the baroreflex in response to an orthostatic maneuver. The fact that it was observed also in the post-AMI patients, together with the observation that none of the groups denoted significant variations of the local causal gain moving from rest to tilt, seems to suggest that the baroreflex response to postural stress is preserved, in terms of increased effectiveness and unaltered sensitivity, after myocardial infarction. On the other hand, the statistically significant decrease of the global causal gain observed with tilt only in AMI patients points out a decreased BRS due to tilt, which agrees with previous findings indicating a reduced capability to cope with the postural challenge after AMI [7,43]. We hypothesize that the symptoms of orthostatic intolerance manifested after AMI are associated with fluctuations of RR and SAP which are not confined within the LF band of the spectrum.

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Along the feedforward direction, a significant decrease of the causal coupling after head-up tilt was detected in Young subjects but not consistently in Old or post-AMI patients. In healthy subjects, a statistically significant decrease with tilt of the coupling from RR to SAP was reported using causal methods in the time domain [49], where it was investigated in terms of the cascade of interactions from RR to diastolic blood pressure (DBP) and then to SBP. Physiologically, this causal coupling is associated with the cardiac run-off, the Windkessel effect and the Frank-Starling law, according to which modifications of the heart period affect the end diastolic volume and then the strength of the systolic contraction [49,50]. The tilt-induced decrease of the effectiveness of the feedforward mechanism can be associated by the increased heart rate which limits the cardiac run-off and consequently the systolic contraction, but other mechanisms cannot be excluded as blood pressure variability is also due to variations in the sympathetic blood vessels control [36]. The lack of a consistent

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570 reduction with tilt of the feedforward coupling in Old and AMI patients could thus be associated to an impairment of these physiological mechanisms related to age and disease.

575 The feedforward gain was found to increase in Old and post-AMI patients after orthostatic stress. Since the occurrence of acute myocardial infarction is thought to be responsible of a significant stiffening of the cardiac muscle, and aging is associated with a stiffening of the vascular bed, these alterations might lead to an alteration of the mechanical effects which allow heart period to drive SAP variability during head-up tilt. A role may be also played by the neural autonomic control, with a larger reduction of heart rate variability compared to SAP variability in AMI patients. The feedforward gain was found to significantly increase with the postural stress also in Old compared to Young subjects. The alteration of the capability of RR to drive SAP variability reported in Old people during tilt suggests that also aging can be responsible for a modification of feedforward mechanisms of mechanical nature that characterize the interactions from the heart period to the systolic arterial pressure. Such results are in agreement with other findings reported in the literature [7] that highlighted an unbalanced RR-SAP regulation in old subjects with increased feedforward and decreased feedback mechanism.

#### 4.4 Limitations and future studies

590 The present study has some limitations that should be taken into account. First of all, according to the study protocol [7], cardiovascular signals of the AMI group were recorded at predischARGE time on patients in pharmacological washout. Only a small subgroup of very low-risk post-AMI patients able to support the suspension of  $\beta$ -blocker treatment without appreciable risk and no taking of antiarrhythmic drugs was involved in the study. Therefore, the results of our study cannot be generalized to the general post-AMI population. Nonetheless, residual effects of the treatment with betablockers may still be present in the AMI patients and thus affect the analyzed cardiovascular dynamics [26].

595 Other limitations are related to the small sample size of Young (19 subjects) and especially Old (12 subjects) groups, and to the fact that such groups are not homogeneous in the gender distribution (males are prevalent in AMI, females are prevalent in Old, while the gender is balanced in Young) [26].

600 A methodological limitation consists in the selection of the order of the parametric model used to fit the time series. Model order selection is an issue in real data analyses, where the true order is usually unknown. A correct model order assessment is rather difficult because the estimated order may not meet the user expectations (in terms of spectral resolution when it is too low, or in terms of interpretability of highly variable spectral profiles when it is too high). In the present study, the use of the Akaike Information Criterion [27] for model order selection led sometimes to spectral contributes of difficult interpretation or even negative power values after spectral decomposition. For this reason, the manual selection of the model order  $p$  could represent a possible workaround to avoid negative power contributes and to obtain a



610 better spectral representation in LF band, still maintaining a tradeoff between good  
data resolution and reasonably low model complexity.

615 Future activities may also include further studies on larger groups of subjects, or on  
patients affected by different pathologies (e.g. hypertension [51,52] or atherosclerosis  
[53,54]). Moreover, an improvement of the automatic order selection algorithm to  
avoid negative power values or the application of other criteria (e.g. Bayesian  
Information Criterion) may be envisaged [55].

## 5. Conclusion

620 This study emphasizes the importance of combining the novel method of spectral  
decomposition [21,56] and a causal approach to cross-spectral analysis [10,13] to  
investigate the coupling and gain mechanisms underlying the closed-loop  
cardiovascular regulation in healthy and diseased states. Combining such approaches  
allows to quantify objectively, at specific well-defined frequencies, the causal  
625 contribution of SAP to RR along the baroreflex pathway and of RR to SAP along the  
mechanical feedforward. The application of the proposed method to cardiovascular  
time series of Young, Old and AMI subjects highlighted that causal local measures  
perform better than traditional non-causal approaches in the evaluation of BRS,  
suggesting that aging and infarction generate impairment of BRS occurring at rest and  
when carrying out the orthostatic maneuver.

630 These findings support the concept that the use of a spectral decomposition  
approach as well as the implementation of causality in the study of interactions  
between heart rate and arterial pressure allows to mitigate the confounding effects of  
other variables operating at different frequencies and of reverse-side mechanisms and  
is thus essential to achieve a more accurate BRS assessment in physio-pathological  
states and different postural conditions.

## 635 CRediT authorship contribution statement

**Riccardo Pernice:** Methodology, Software, Visualization, Writing-Original draft  
preparation. **Laura Sparacino:** Software, Visualization, Writing-Original draft  
preparation. **Giandomenico Nollo:** Data curation. **Alessandro Busacca:** Writing-  
640 Reviewing and Editing. **Luca Faes:** Methodology, Formal analysis, Validation,  
Writing- Reviewing and Editing, Supervision

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