

Comparison of discretization strategies for the model-free information-theoretic assessment of short-term physiological interactions

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This work presents a comparison between different approaches for the model-free estimation of information-theoretic measures of the dynamic coupling between short realizations of random processes. The measures considered are the Mutual Information Rate (MIR) between two random processes X and Y and the terms of its decomposition evidencing either the individual entropy rates of X and Y and their joint entropy rate, or the transfer entropies from X to Y and from Y to X and the instantaneous information shared by X and Y . All measures are estimated through discretization of the random variables forming the processes, performed either via uniform quantization (binning approach) or rank ordering (permutation approach). The binning and permutation approaches are compared on three datasets including short realizations of cardiorespiratory (CR, heart period and respiration flow), cardiovascular (CV, heart period and systolic arterial pressure) and cerebrovascular (CB, mean arterial pressure and cerebral blood flow velocity) measured in different physiological conditions, i.e., spontaneous vs. paced breathing or supine vs. upright positions. Our results show that, with careful selection of the estimation parameters (i.e., embedding dimension and number of quantization levels for the binning approach), physiologically meaningful patterns of the MIR and of its components can be achieved in the analyzed datasets. We found that paced breathing at slow breathing rates induces less complex and more coupled CR dynamics, while postural stress leads to an unbalancing of CV interactions with prevalent baroreflex coupling and to less complex pressure dynamics with preserved CB interactions. These results are better highlighted by the permutation approach, thanks to its more parsimonious representation of the discretized dynamic patterns, which allows to explore interactions with longer memory while limiting the curse of dimensionality.

Model-free tools for the analysis of bivariate time series are fundamental to provide a proper description of how the coupling among systems arises from the underlying possibly non-linear regulatory mechanisms. Among the tools devised to perform this analysis, information-theoretic measures are largely explored to infer short-term interactions from pairs of physiological time series. In this context, the so-called mutual information rate (MIR) is a long-known measure of the dynamic interaction between two random processes, which can be reliably approximated from bivariate linear models fitting the observed time series, but is much harder to quantify directly if the model assumptions are not satisfied. In this work, the challenge of performing model-free estimation of this measure is faced exploiting the possibility to decompose the MIR into terms evidencing measures of entropy rate and conditional mutual information, executing low-dimensional time series embedding, and implementing the estimation of these terms through binning- and permutation-based discretization strategies. While the application to short-term cardiorespiratory, cardiovascular and cerebrovascular variability series provides physiologically plausible results, it also evidences troublesome aspects that call for the development of improved entropy estimators and refined embedding strategies.

I. INTRODUCTION

In recent years, a wide range of approaches have been proposed to assess the temporal statistical structure and the interaction between coupled dynamic processes from the analysis of bivariate time series taken as realizations of these processes. The most prominent methods are based on state space interdependence¹⁻³, correlation analyses⁴⁻⁶, or on the concept of Granger causality (GC) implemented in the time or frequency domains⁷⁻⁹. A general and flexible framework, which encompasses most of these approaches, is the framework of information dynamics^{10,11}, which has been developed and widely exploited to characterize the interdependence of coupled systems in several fields including neuroscience^{12,13} and physiology¹⁴⁻¹⁶. In particular, entropy-based measures quantifying the dynamical complexity or information storage within a single process^{17,18}, or the directed information transfer from one process to another^{19,20}, have been successfully used to assess physiological interactions in the cardiorespiratory^{15,21}, cardiovascular^{22,23} and cerebrovascular^{24,25} systems.

A long-known and well-defined measure to assess the dynamic coupling between two time series using information theory is the Mutual Information Rate (MIR), which quantifies the information exchanged over time by two coupled systems^{26,27}. However, since the MIR is defined for infinite-

dimensional variables, its reliable estimation in practical contexts where short stationary realizations are typically available is a daunting task. In spite of this, a renewed interest for this measure has recently emerged thanks to the possibility to decompose it in different forms evidencing measures of entropy rate or conditional mutual information²⁸ and to the development of accurate estimators for the constituent terms^{29–32}. The interest in this approach stems also from the fact that the terms appearing in the decomposition of the MIR constitute popular measures such as the conditional entropy³³ or the transfer entropy¹⁹, which are related respectively to the concepts of complexity and causality, whose computation is relevant *per se* in several applications. Nevertheless, the analysis of the MIR and its constituent terms has been faced up to now for specific classes of coupled random processes, such as continuous-time point processes³⁰ or discrete-time linear Gaussian processes³⁴. The estimation of the MIR and its composing terms in the more general case where assumptions about the type of processes or about the form of interactions to be analyzed are relaxed is a challenging task that has received less attention up to now.

The present work faces the problem of estimating the MIR and the terms of its decomposition using model-free entropy estimates. Model-free approaches are alternative to parametric estimators, which have the advantage of favoring computation but the drawback of losing generality. Indeed, regardless of whether linear³⁵ or nonlinear⁷ model-based methods are used, constraining the analysis on a specific model structure limits the capability to detect and describe complex interactions and multifaceted dynamic behaviors. Here, we employ a model-free strategy for the computation of the information shared dynamically between two random processes, which is described as follows. First we express the MIR equivalently as the sum of the individual entropy rates of each process minus their joint entropy rate, and as the sum of the transfer entropy computed along the two directions of interaction plus the instantaneous information shared by the processes at zero lag. Second, to allow computation of the measures from finite-length realizations of the processes, we perform low-dimensional embedding approximating the past histories of the processes with a small number of time-lagged components. Third, among the variety of methods for the non-parametric estimation of dynamic information measures³⁶, we employ discretization methods, which compute the relevant entropy terms on discrete random variables obtained through symbolization of the observed continuous-valued vector variables. Specifically, the approaches to discretization used herein are the binning method performing uniform quantization of the time series samples³⁷ and the permutation method working on the ranks of the amplitude values within patterns extracted from the time series³⁸.

The binning- and permutation-based approaches for discretization are compared extensively in three datasets of physiological time series collected from young healthy subjects and including different pairs of series measured in different experimental protocols: (i) heart period and respiratory flow measured during spontaneous and paced breathing, (ii) heart period and systolic arterial pressure measured at rest and dur-

ing head-up tilt, and (iii) mean arterial pressure and cerebral blood flow velocity measured at rest and during head-up tilt. These applications have been chosen because they feature challenging conditions where complex short-term physiological mechanisms need to be inferred using model-free methods applied on short realizations (~ 250 -300 points). In this context, the comparison between the two discretization approaches is made observing their ability to reveal expected physiological responses while facing two contrasting needs, i.e., working in low-dimensional spaces and providing an accurate description of the process dynamics in terms of time-lagged interactions and signal amplitude.

II. METHODS

Let us consider two possibly interacting dynamical systems \mathcal{X} and \mathcal{Y} , and assume that their joint evolution over time is described by the stochastic processes $X = \{X_n\}$ and $Y = \{Y_n\}$, where n is the time counter. To introduce the notation, let us indicate as X_n, Y_n the scalar random variables describing the current state of X and Y , as $X_{n-k:n-1} = [X_{n-k} \cdots X_{n-1}]$ and $Y_{n-k:n-1} = [Y_{n-k} \cdots Y_{n-1}]$ the k -dimensional vector variables sampling X and Y over the past k lags, and as $X_{<n} = \lim_{k \rightarrow \infty} (X_{n-k:n-1})$ and $Y_{<n} = \lim_{k \rightarrow \infty} (Y_{n-k:n-1})$ the infinite-dimensional variables sampling X and Y over their whole past history. In this work we consider the problem of estimating the dynamic coupling between the two systems through the so-called Mutual Information Rate (MIR)²⁶. In the following subsections we will first provide the theoretical formulation of the MIR and of its information-theoretic decomposition, and then illustrate the binning and permutation approaches for the estimation of the MIR and its terms starting from two scalar time series $\mathbf{x} = \{x_n\}_{n=1}^N$ and $\mathbf{y} = \{y_n\}_{n=1}^N$, observed as realizations of the processes X and Y under analysis.

A. Mutual Information Rate Decomposition

The MIR between two stationary and ergodic stochastic processes X and Y is defined as

$$I_{X;Y} = \lim_{k \rightarrow \infty} \frac{1}{k} I(X_{n-k:n-1}; Y_{n-k:n-1}), \quad (1)$$

where $I(\cdot; \cdot)$ is the mutual information between two random variables. The MIR is a dynamic measure of the information exchanged per unit of time between two dynamical systems²⁶, which has been implemented in different forms to quantify dynamic interactions between physiological processes^{29,30,32}. The popularity of this measure stems also from the fact that it can be decomposed evidencing information quantities that have meaningful interpretation and are practically computable from time series data. In fact, starting from the equivalent definitions of entropy rate of an ergodic stochastic process X ³⁹:

$$H_X = \lim_{k \rightarrow \infty} \frac{1}{k} H(X_{n-k:n-1}) = H(X_n | X_{<n}), \quad (2)$$

where $H(\cdot)$ denotes entropy and $H(\cdot|\cdot)$ denotes conditional entropy, some elaborations (see, e.g., the supplemental material of Mijatovic et al.³⁰) lead to the expansions:

$$I_{X;Y} = H_X + H_Y - H_{X,Y}, \quad (3)$$

$$I_{X;Y} = T_{X \rightarrow Y} + T_{Y \rightarrow X} + I_{X,Y}. \quad (4)$$

The first decomposition of the MIR given in (3) evidences how it can be formulated comparing the sum of the individual entropy rates of the analyzed processes X and Y , H_X and H_Y , with their joint entropy rate $H_{X,Y}$; the entropy rate is a well-known measure of complexity expressed as the conditional entropy of the present state of a process given its own past history^{37,40}. On the other hand, the decomposition (4) expresses the MIR as the sum of the two transfer entropies from X to Y and from Y to X , $T_{X \rightarrow Y}$ and $T_{Y \rightarrow X}$, plus a term quantifying the instantaneous information shared by X and Y at zero lag, $I_{X,Y}$; the transfer entropy (TE) is a very popular measure of directed information transfer related to the concept of Granger causality¹⁹, while the instantaneous transfer is a symmetric measure related to the concept of instantaneous causality²⁸.

The two decomposition of the MIR are illustrated graphically in Fig. 1 making use of Venn diagrams. This representation helps understanding how all the measures appearing in the decompositions (3) and (4) can be expressed in terms of conditional entropies, i.e.:

$$H_X = H(X_n|X_{<n}); \quad (5a)$$

$$H_Y = H(Y_n|Y_{<n}); \quad (5b)$$

$$H_{X,Y} = H(X_n, Y_n|X_{<n}, Y_{<n}); \quad (5c)$$

$$T_{X \rightarrow Y} = H(Y_n|Y_{<n}) - H(Y_n|X_{<n}, Y_{<n}); \quad (5d)$$

$$T_{Y \rightarrow X} = H(X_n|X_{<n}) - H(X_n|X_{<n}, Y_{<n}); \quad (5e)$$

$$I_{X,Y} = H(X_n|X_{<n}, Y_{<n}) - H(X_n|X_{<n}, Y_n, Y_{<n}). \quad (5f)$$

The expressions in (5) constitute the basis for the practical computation of the MIR decompositions adopted in this work. First, under the Markov assumption stated with memory m , the past history of each process is approximated with m -dimensional vector variables, i.e., $X_{<n} \approx X_{n-m:n-1}$ and $Y_{<n} \approx Y_{n-m:n-1}$ (note that $[Y_n Y_{<n}]$ in (5f) becomes $Y_{n-m:n}$). Moreover, we exploit the fact that any conditional entropy can be expressed as the difference between two entropy terms: for instance, the approximation of the conditional entropy in (5c) can be written as $H(X_n, Y_n|X_{n-m:n-1}, Y_{n-m:n-1}) = H(X_{n-m:n}, Y_{n-m:n}) - H(X_{n-m:n-1}, Y_{n-m:n-1})$. Therefore, computation amounts to estimate the entropy of the vector variables sampling the present and past states of the processes combined as described above, and then to plug this estimates into (5) and (3,4) to obtain estimates of the MIR and of its terms. In the next subsections we present two symbolization methods for the computation of the MIR decomposition, which estimate the entropy of discrete random variables obtained from the observed continuous variables either using quantization levels or ordinal patterns.

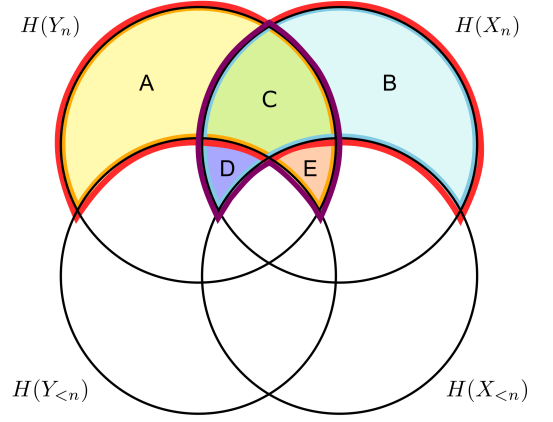


FIG. 1. Venn diagram depicting the MIR decomposition measures defined for two random processes X and Y . Taking the areas of the four circles as the entropy of the present state (upper circles) and of the past history (lower circles) of X (right) and Y (left) one can see that: the entropy rate of X is the union of the portions B, C and D (H_X , Eq. (5a), area with azure contour); the entropy rate of Y is the union of A, C and E (H_Y , Eq. (5b), area with yellow contour); the entropy rate of $\{X, Y\}$ is the union of A, B and C ($H_{X,Y}$, Eq. (5c), area with red contour); the transfer entropy from X to Y is the portion E ($T_{X \rightarrow Y}$, Eq. (5d), orange area); the transfer entropy from Y to X is the portion D ($T_{Y \rightarrow X}$, Eq. (5e), blue area); and the instantaneous information shared between X and Y is the portion C ($I_{X,Y}$, Eq. (5f), green area). As a result, the MIR between X and Y is the union of C, D and E ($I_{X;Y}$, Eqs. (3,4), area with purple contour).

B. Binning approach

The binning method is probably the most intuitive approach to estimate entropy measures for continuous random variables via symbolization^{36,37}. This approach is based on building discrete random variables through quantization of the continuous-valued variables and then computing entropies from the probabilities estimated for the discrete variables via the frequentistic approach. For a generic random variable V taking values in the continuous domain $\mathcal{D}_V = [v_{\min}, v_{\max}]$, quantization returns a discrete random variable B_V taking values in the alphabet $\mathcal{A}_{B_V} = \{1, \dots, b\}$ formed by b quantization levels, or bins; in the case of uniform quantization, the discretization rule is that an observation of V , $v \in \mathcal{D}_V$, is transformed in the observation of B_V , $b_v = i$, if $v_{\min} + (i-1)r \leq v < v_{\min} + ir$, where $r = (v_{\max} - v_{\min})/b$ is the bin amplitude. Once binning is performed, the probability of each discrete value in \mathcal{A}_{B_V} is estimated naturally as the frequency of occurrence of the value over many observations, and an estimate of the entropy of V is the entropy of B_V obtained by the classical formula

$$\hat{H}(V) = - \sum_{b_v \in \mathcal{A}_{B_V}} \hat{p}(b_v) \log \hat{p}(b_v). \quad (6)$$

The formulations above hold intuitively also for vector variables, for which quantization is applied over each scalar component; in the case of an original d -dimensional continuous variable $\mathbf{V} = [V_1 \dots V_d]$, which is quantized using b bins

for each dimension, the corresponding discrete vector random variable $B_{\mathbf{V}} = [b_{V_1} \cdots b_{V_d}]$ takes values inside an alphabet formed by b^d symbols.

In the case of two random processes X and Y relevant for this work, the vector random variables whose entropies are needed to compute the measures listed in (5) are formed combining the present state and the histories sampled up to m past lags of the two processes. For instance, the binning estimator of the conditional entropy term H_X in (5a) works with the m -dimensional variable $X_{n-m:n-1}$ sampling the past of X and with the $(m+1)$ -dimensional variable $X_{n-m:n} = [X_{n-m:n-1} X_n]$ sampling the present and the past of X . Starting from the time series $\mathbf{x} = \{x_n\}_{n=1}^N$, $N-m$ observations of $X_{n-m:n-1}$ and $X_{n-m:n}$ are obtained which take, after quantization, values inside alphabets composed by b^m and b^{m+1} symbols, respectively. Then, after estimating the probability of occurrence of each observation, the entropies $\hat{H}(X_{n-m:n-1})$ and $\hat{H}(X_{n-m:n})$ are estimated using (6), and the entropy rate is computed as $\hat{H}_X = \hat{H}(X_{n-m:n}) - \hat{H}(X_{n-m:n-1})$. The same formulation can be applied for estimating the entropy rate H_Y starting from the time series $\mathbf{y} = \{y_n\}_{n=1}^N$, while the joint vector variables $[X_{n-m:n-1} Y_{n-m:n-1}]$ and $[X_{n-m:n} Y_{n-m:n}]$ are needed to compute the joint entropy rate (5c); in the latter case, the discrete variables visit spaces of dimension b^{2m} and b^{2m+2} , and need to be estimated from $N-m$ observations. The computation of the TE and instantaneous information sharing in Eqs. (5d,e,f) is even more intricate since it requires the computation of four entropy terms, as each of the two conditional entropy terms is defined as the difference of entropy terms according to the rule $H(V|W) = H(V,W) - H(W)$ for any variables V and W . In the case of the TE, they involve the vector variables sampling the past history of the target process taken alone (dimension b^m) or together with the present state of the same process (dimension b^{m+1}) and the past of both processes taken alone or together with the present of the target (dimensions b^{2m} , b^{2m+1}); in the case of the instantaneous information shared by the two processes, the four entropies involve the history of both processes taken alone (dimension b^{2m}), with the present of one of the two processes (dimension b^{2m+1}), or with the present of both processes (dimension b^{2m+2}).

C. Permutation-based approach

The approaches based on permutations perform symbolization working directly on vector variables, taking into account the amplitude order of neighboring samples within the realizations of these variables without effective consideration of their absolute amplitude values³⁸. For a generic d -dimensional continuous random variable $\mathbf{V} = [V_1 \cdots V_d]$, realizations of the associated discrete variable $R_{\mathbf{V}}$ are obtained through a rank ordering procedure as follows: if $\mathbf{v} = [v_1 \cdots v_d]$ is a realization of \mathbf{V} , the corresponding realization of $R_{\mathbf{V}}$ is $r_{\mathbf{v}} = [r_{v_1} \cdots r_{v_d}] \in \mathcal{A}_{R_{\mathbf{V}}}$, where $r_{v_i} \in \{1, \dots, d\}$ is the rank order of v_i inside the sequence \mathbf{v} rearranged in ascending order (e.g., $r_{v_i} = 1$ if $v_i = \min(\mathbf{v})$ and $r_{v_i} = d$ if $v_i = \max(\mathbf{v})$); for two equal components of \mathbf{v} the smallest rank is assigned to the component appearing last). Once discretization is performed,

the probability of each discrete vector value belonging to the alphabet $\mathcal{A}_{R_{\mathbf{V}}}$ is estimated naturally as the frequency of occurrence of the value over many observations, and an estimate of the entropy of \mathbf{V} is the entropy of $R_{\mathbf{V}}$ obtained as

$$\hat{H}(\mathbf{V}) = - \sum_{r_{\mathbf{v}} \in \mathcal{A}_{R_{\mathbf{V}}}} \hat{p}(r_{\mathbf{v}}) \log \hat{p}(r_{\mathbf{v}}). \quad (7)$$

We note that the discrete random variable $R_{\mathbf{V}}$ obtained applying the permutation strategy to the continuous d -dimensional variable \mathbf{V} takes values inside an alphabet with cardinality $|\mathcal{A}_{R_{\mathbf{V}}}| = d!$, which is usually smaller than the cardinality of the alphabet obtained quantizing the variable with b bins, $|\mathcal{A}_{B_{\mathbf{V}}}| = b^d$. This favors the permutation strategy for the estimation of entropy from a limited number of observations of the variable under analysis.

As for the binning approach, the entropy measures needed to compute the terms of the MIR decomposition in (5) are estimated from observations of the present and past states of the analyzed processes X and Y . Again, the variables $X_{n-m:n}$, $X_{n-m:n-1}$ and $Y_{n-m:n}$, $Y_{n-m:n-1}$ are considered for the estimation of the individual entropy rates \hat{H}_X and \hat{H}_Y in (5a,b), and the joint variables $[X_{n-m:n} Y_{n-m:n}]$ and $[X_{n-m:n-1} Y_{n-m:n-1}]$ are considered to estimate the joint entropy rate $\hat{H}_{X,Y}$ in (5c); the alphabet sizes are $(m+1)!$ and $m!$, and $((m+1)!)^2$ and $(m!)^2$, respectively for the individual and joint entropy rates. As regards the computation of the TE terms in (5d,e), the involved variables are the past history of the target process taken alone or together with its present state (alphabet sizes $m!$ or $(m+1)!$), and the history of both processes taken alone or together with the past of the target (alphabet sizes $(m!)^2$ or $m! \cdot (m+1)!$). Finally, the variables involved in the computation of the instantaneous information sharing are those covering the history of both processes taken alone, taken with the present of one of the two processes, or taken with the present of both processes, which have alphabet size $(m!)^2$, $m! \cdot (m+1)!$, and $((m+1)!)^2$, respectively.

It is worth noting that the approach which we follow to compute any measure of entropy rate is different than that used in several studies working with permutations where the observable corresponding to the present state of a system (e.g., X_n) is converted to a rank vector^{41,42}, and is rather similar to that followed in the works of Kugiumtzis^{43,44} where such an observable is taken as a scalar and is considered together with the observable of the past m states of the same system (e.g., $X_{n-m:n-1}$) in the formation of the joint state (e.g., $X_{n-m:n}$) from which rank ordering is performed. This approach conforms with the definition of conditional entropy, and has been shown to lead to less biased entropy estimates⁴³.

III. APPLICATIONS TO PHYSIOLOGICAL TIME SERIES

This section reports the application of the two strategies presented in Sect. II for the computation of the dynamic information measures composing the MIR to three different types of pairwise interactions typically analyzed in the context of short-term physiological variability. Specifically, the proposed methodology is applied to assess cardiorespiratory

interactions during spontaneous and paced breathing (Section III A.1), as well as cardiovascular or cerebrovascular interactions during rest and postural stress (Sections III A.2 and III A.3).

A. Experimental protocols

1. Cardiorespiratory variability analysis

Cardiorespiratory interactions were investigated exploiting an historical database collected for the analysis of short-term physiological variability during paced breathing⁴⁵ (*database 1*). The database includes physiological time series measured from 19 young healthy subjects monitored in the resting supine position during four experimental conditions, i.e., spontaneous breathing (SB) and controlled breathing at 10 breaths/minute (C10), 15 breaths/minute (C15) and 20 breaths/minute (C20). The one time series is the sequence of the heart periods (RR intervals) extracted from the electrocardiographic (ECG) signal, given as the time intervals between two consecutive ECG R peaks. The other time series contains the respiratory amplitude values (RESP) extracted from the nasal respiration flow signal sampled at the onset of each heart period (see Fig. 2a). For each subject and condition, synchronous time series of 256 values were considered for the analysis. Further details on the experimental protocol, signal acquisition, and time series extraction procedure can be found in the reference paper⁴⁵.

2. Cardiovascular variability analysis

Cardiovascular interactions were investigated exploiting a historical dataset collected for the study of the short-term physiological response of the cardiovascular system to postural stress⁴⁶ (*database 2*). The database includes physiological time series measured from 15 young healthy subjects monitored in the resting supine position (R) and in the 60° upright position (T) reached after a passive head-up tilt manoeuvre. The one time series contains 300 beats of RR intervals defined as above. The other time series contains the systolic arterial pressure (SAP), given as the local maxima of the the continuous photoplethysmographic arterial pressure signal (volume-clamp method) measured within each detected RR interval (see Fig. 2b). Further details on the experimental protocol, signal acquisition, and time series extraction procedure can be found in the reference paper⁴⁶.

3. Cerebrovascular variability analysis

Cerebrovascular interactions were investigated exploiting a historical database collected to study the response of cerebral autoregulation to postural stress²⁵ (*database 3*). The database includes physiological time series measured from 13 young healthy subjects monitored in the resting supine position (R) and in the 60° upright position (T) reached after

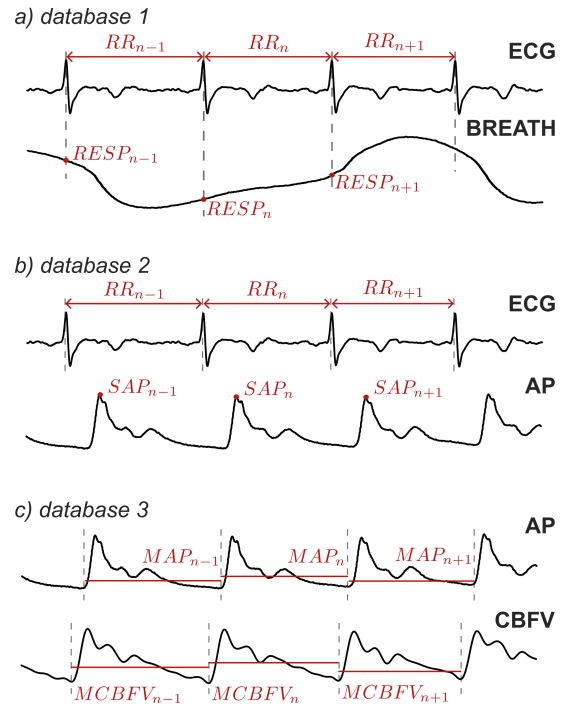


FIG. 2. Schematic representation of the procedure for extracting the time series from the physiological signals synchronously acquired in the three applications. (a) Cardiorespiratory variability analysis (*database 1*): the n^{th} RR value is the time interval between the n^{th} and the $(n+1)^{\text{th}}$ R peaks of the electrocardiogram (ECG), while the n^{th} RESP value is obtained sampling the respiration airflow signal (BREATH) in correspondence of the n^{th} R peak. (b) Cardiovascular variability analysis (*database 2*): the n^{th} SAP value is taken as the maximum value of the arterial pressure (AP) signal measured within the n^{th} RR interval. (c) Cerebrovascular variability analysis (*database 3*): from the arterial pressure (AP) and cerebral blood flow velocity (CBFV) signals, the n^{th} values of the mean AP (MAP) and mean CBFV (MCBFV) are obtained by averaging each signal within the n^{th} detected diastolic pulse interval defined as the time interval between two consecutive local minima.

a passive head-up tilt manoeuvre. Here, we analyzed time series of 250 beats of the mean arterial pressure (MAP) and mean cerebral blood flow velocity (MCBFV), obtained from the continuous photoplethysmographic arterial pressure signal (volume-clamp method) and cerebral blood flow velocity signal (transcranial doppler method) respectively as the mean signal values measured between each pair of consecutive diastolic points (i.e., local minima, see Fig. 2c). Further details on the experimental protocol, signals acquisition, and time series extraction procedure can be found in the reference papers²⁵.

B. Data analysis

The dynamic information measures composing the MIR were computed on the pairs of time series measured as described in the previous subsection, which were interpreted as realizations of the random processes $\{X, Y\}$ descriptive of car-

diorespiratory interactions ($X = \text{RR}, Y = \text{RESP}$, *database 1*), cardiovascular interactions ($X = \text{RR}, Y = \text{SAP}$, *database 2*), and cerebrovascular interactions ($X = \text{MAP}, Y = \text{MCBFV}$, *database 3*). The two discretization approaches presented in Sect. II were employed to decompose the MIR either as the sum of the individual entropy rates of the two processes (i.e., H_X and H_Y) minus their joint entropy rate (i.e., $H_{X,Y}$), or as the sum of the two transfer entropies from one process to another (i.e., $T_{X \rightarrow Y}$ and $T_{Y \rightarrow X}$) plus the instantaneous information term (i.e., $I_{X,Y}$).

1. Parameter setting

A crucial aspect in the implementation of the discretization approaches is the selection of the free parameters for the analysis, i.e., the length of the memory used to cover the past history of the processes (parameter m) and, for the binning estimator, the number of quantization levels used for coarse graining (parameter b). The choice of these parameters is connected to the issue of estimating the entropy of high-dimensional variables from finite-size datasets, which is related to the so-called curse of dimensionality^{37,47}. In the practical application of symbolization methods, empirical criteria suggest to optimize the parameters in a way such that the alphabet size remains as low as the length N of the series³⁷. Considering the most challenging condition in which the past and present states of both processes needs to be covered to estimate the entropy of the vector variable $[X_{n-m:n} Y_{n-m:n}]$, the alphabet size is equal to b^{2m+2} in the case of binning and to $((m+1)!)^2$ for the permutation approach. Accordingly, we set $b = 4$ and $m = 1$ when the binning approach was implemented, so as to deal with a number of quantization levels equal to $4^{2 \cdot 1 + 2} = 256$ and adhere with the empirical rule stated above; we note that these values are lower than those typically used for the complexity analysis of individual time series (i.e., $m = 2, b = 6$)³⁷. As regards the permutation approach, we used the pattern length typically adopted in permutation entropy analyses, which is the minimum advised to guarantee variability in the discrete patterns^{38,48}, i.e. $m = 3$. However, as this choice corresponds to a maximum alphabet size of $((3+1)!)^2 = 576$, we expect that the estimates involving the higher-dimensional variables can exhibit non-negligible bias.

2. Surrogate data analysis

The statistical significance of the computed information measures was assessed at a single-subject level via the use of surrogate data generated specifically for each measure. In particular, to assess the significance of conditional entropy measures (i.e., H_X , H_Y and $H_{X,Y}$) there is the need to destroy the temporal structure. Therefore, random shuffle surrogates⁴⁹ were generated according to the null hypothesis of independent and identically distributed random variables, permuting randomly and independently the order of the samples in the two series. On the contrary, to assess the significance of conditional mutual information measures (i.e., $T_{X \rightarrow Y}$, $T_{Y \rightarrow X}$, and

$I_{X,Y}$), as well as of the MIR, it is sufficient to destroy the coupling of the two series, while it is preferable to maintain the statistical properties of the individual series⁵⁰. Therefore, random time shift bivariate surrogates were generated according to the null hypothesis of independent random processes, shifting the samples of the time series Y over time (while wrapping the extra values around the beginning of the series) and leaving the other series unchanged³; the shift was chosen randomly, imposing a minimum shift of 20 lags.

For each pair of original time series, 100 pairs of surrogate time series were generated and the considered information measure was computed both on the original series and on the 100 surrogate pairs. Then, a non-parametric test based on percentiles was adopted, distinguishing the two type of measures: each conditional entropy measure was deemed as statistically significant if its value computed on the original series was lower than the 5th percentile of its distribution derived from the random shuffling surrogates; each conditional mutual information measure was deemed as statistically significant if its value computed on the original series was higher than the 95th percentile of its distribution derived from the random time shift surrogates.

3. Statistical analysis

In order to assess the statistical significance of the variations of a given measures among different conditions, the following parametric tests were applied. As regards cardiorespiratory interactions, one-way analysis of variance (ANOVA) was applied, followed by a post-hoc paired Student's t-test with Bonferroni correction for multiple comparisons ($n=3$: SB vs. C10, SB vs. C15, SB vs. C20). As regards cardiovascular and cerebrovascular analyses, the paired Student's t-test was directly carried out between the two conditions under test. The use of parametric tests is justified by the fact that normality of the distributions was verified for all measures and all three databases through one-sample Kolmogorov-Smirnov test. All statistical tests were performed at a 5% significance level.

Moreover, to assess the magnitude of the effect that a change in the experimental condition has on each analyzed index, the effect size was assessed using the Cohen's d index⁵¹. The index quantifies the difference between the means of the two distributions under analysis divided by the pooled standard deviation

$$d = \frac{\bar{x}_1 - \bar{x}_2}{\sqrt{\frac{(n_1-1)s_1^2 + (n_2-1)s_2^2}{n_1+n_2-2}}} \quad (8)$$

being \bar{x} , s , and n the mean, the standard deviation, and the number of samples (i.e., the number of subjects) of the two distributions under comparison, respectively. Usually, the effect size is considered small, medium, and large, if the absolute value of d is lower than 0.2, between 0.2 and 0.5, or higher than 0.8, respectively.

C. Results and Physiological Interpretation

1. MIR analysis

Fig. 3 reports the distribution of the MIR values estimated for the three datasets described above using the binning approach (a-c) and the permutation approach (d-f) to discretize the observed continuous-valued time series.

As regards cardiorespiratory interactions, for both approaches the ANOVA test highlighted significant variations across the phases of the breathing protocol. As reported in Table I, the p -value of the ANOVA test is far lower in the case of binning, for which the post-hoc paired tests evidence significantly higher values of $I_{RR;RESP}$ during paced breathing at 10 and 15 breaths/minute than during spontaneous breathing (Fig. 3a), also in association with a large effect size (Cohen's $|d| > 0.8$). A different trend is documented using the permutation method, for which a non-significant tendency to lower MIR values during C15 and to higher MIR values during C20 is observed (Fig. 3d). Generally, the estimates obtained through the binning approach are significant for a larger number of subjects, although permutation-based method returns higher values.

With regard to cardiovascular interactions (Figs. 3b,e), no statistically significant differences in the MIR between RR interval and systolic pressure series are observed comparing the orthostatic stress and supine rest conditions. The estimates of $I_{RR;SAP}$ tend to exhibit opposite responses to the postural change, with tendency toward higher values using binning and toward lower values using permutations, both with a medium effect size (Table II).

Similar observations can be made for the cerebrovascular application, with absence of statistically significant differences in the MIR between mean cerebral flow velocity and arterial pressure series (Figs. 3c,d) and limited effect sizes (Table III) observed using both estimators.

Overall, these results suggest that the MIR measure is not very informative when applied to these datasets, as it does not discriminate the different physiological conditions and/or provides indications depending on the estimator adopted. In the following sections we will elaborate more on these aspects, decomposing the MIR in terms of complexity or causality measures and analyzing these measures in the three proposed physiological applications.

2. MIR decomposition: cardiorespiratory variability analysis

Fig. 4 reports the distribution of the information measures composing the MIR computed for the RR and RESP time series using the binning (a-f) and permutation-based (g-l) approaches in the four phases of the paced breathing protocol.

Considering the decomposition evidencing the entropy rate measures computed via conditional entropy, the binning estimator reveals a tendency towards an increase of H_{RR} during paced breathing (Fig. 4a) - with p -value of the ANOVA test slightly above the significance threshold (Table I), the presence of stable values of H_{RESP} across conditions (Fig. 4b),

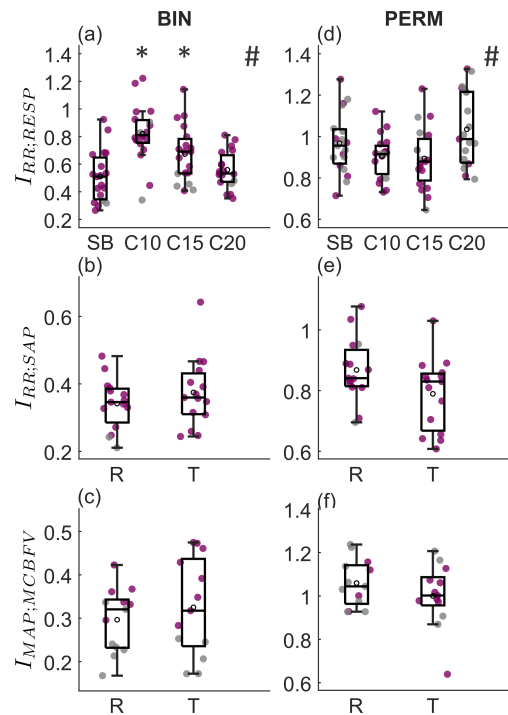


FIG. 3. Analysis of the MIR measure computed on pairs of physiological variability series. Panels report the boxplot distributions and individual values of the MIR computed: (a,d) on cardiorespiratory series (RR, RESP) analyzed during spontaneous breathing (SB) and controlled breathing at 10, 15, 20 breaths/minute (C10, C15, C20); (b,e) on cardiovascular series (RR, SAP) analyzed at rest (R) and after head-up tilt (T); (c,f) on cerebrovascular series (MAP, MCBFV) analyzed at rest (R) and after head-up tilt (T). The analysis was performed using the binning approach (a,b,c) or the permutation approach (d,e,f) for discretization. In each panel, colored and gray circles correspond to values deemed respectively as statistically significant and non-significant using surrogate data analysis; black open circles refer to the sample mean of each distribution. Statistically significant differences, cardiorespiratory series: #, $p < 0.05$, ANOVA; *, $p < 0.05/3$ SB vs. C10/C15/C20, paired t-test with Bonferroni correction; *, $p < 0.05$, R vs. T, paired t-test.

and the presence of lower values during C10 and higher values during C20 compared to SB for the joint entropy rate $H_{RR;RESP}$ (Fig. 4c). On the contrary, the permutation approach evidenced more consistently, when comparing the paced breathing conditions with SB, a significant decrease of all entropy rates during C10 and C15, and comparable or even higher values during C20 (Fig. 4g-i); the effect sizes are also larger when the entropy rate variations are assessed by permutations (Table I).

Physiologically, the lower entropy rates detected by the permutation method during C10 and C15, which can be taken as a marker of less complex RR and RESP dynamics, can be related to the mechanism of respiratory sinus arrhythmia (RSA), i.e., the variation of heart rate due to respiration; this mechanism is enhanced during forced ventilation at low breathing rates⁵², while its effects compared with spontaneous breathing tend to be less pronounced when the respiratory rate

increases⁵³. Moreover, RSA and other cardiorespiratory control mechanisms, e.g., the cardio-ventilatory coupling and the respiratory stroke volume synchronization⁵⁴, are expected to elicit changes in the complexity of cardiac dynamics, associated with similar variations in the complexity of respiratory dynamics; we document that such changes are present as well when RR and RESP time series dynamics are evaluated together. On the other hand, we ascribe the lack of consistent changes in the entropy rate when the binning method is used to the limited exploitation of past values of the RR and RESP series in the evaluation of complexity ($m = 1$, compared with $m = 3$ used for the permutation method).

The different history length used for binning and permutation mentioned above plays probably a role also in the different trends observed considering the decomposition of the MIR between RR and RESP which evidences the causality measures computed via conditional mutual information. For this decomposition, the binning approach evidences higher transfer entropies evaluated in both directions of interaction during C10 and C15 (Fig. 4d,e) along with stable values of the instantaneous information shared between RR and RESP (Fig. 4f). On the contrary, the permutation approach evidences rather stable values of the information transfer (see Fig. 4j,k, where only an increase of $T_{RESP \rightarrow RR}$ during C20 is detected) along with a decrease during C10 and C15 of the information shared instantaneously (see Fig. 4l, showing also that $I_{RR-RESP}$ is barely significant). The patterns of information transfer evidenced by the permutation method are more consistent with those observed on the same dataset using both parametric and model-free estimators^{55,56}, and support the hypothesis that significant physiological changes in cardiac autonomic activity do not occur during supine paced breathing⁵². As regards the instantaneous information shared between RR and RESP, it should be noticed that its variations should be ascribed to directed effects from RESP to RR due to the way the RR and RESP time series were extracted from the physiological signals^{55,56} (see Fig. 2).

3. MIR decomposition: cardiovascular variability analysis

Fig. 5 reports the distribution of the information measures composing the MIR computed for the RR and SAP time series measured at rest and during head-up tilt using the binning (a-f) and permutation-based (g-l) approaches.

In this application, the binning and permutation approaches for the discretization of the observed time series lead to similar results, as documented in Fig. 5. The MIR decomposition computed via conditional entropy shows how the orthostatic stress dampens the complexity of cardiovascular dynamics, as documented by the statistically significant decrease of the joint entropy rate of RR and SAP (Fig. 5c,i). The decrease is mainly driven by the simplification of the cardiac dynamics evidenced by the drop of the entropy rate of RR (Fig. 5a,g), in the presence of an unchanged entropy rate of SAP (Fig. 5b,h). These trends are detected more evidently using the binning estimator, which yields lower p -values and higher effect sizes in the comparison between the rest and tilt conditions (Table

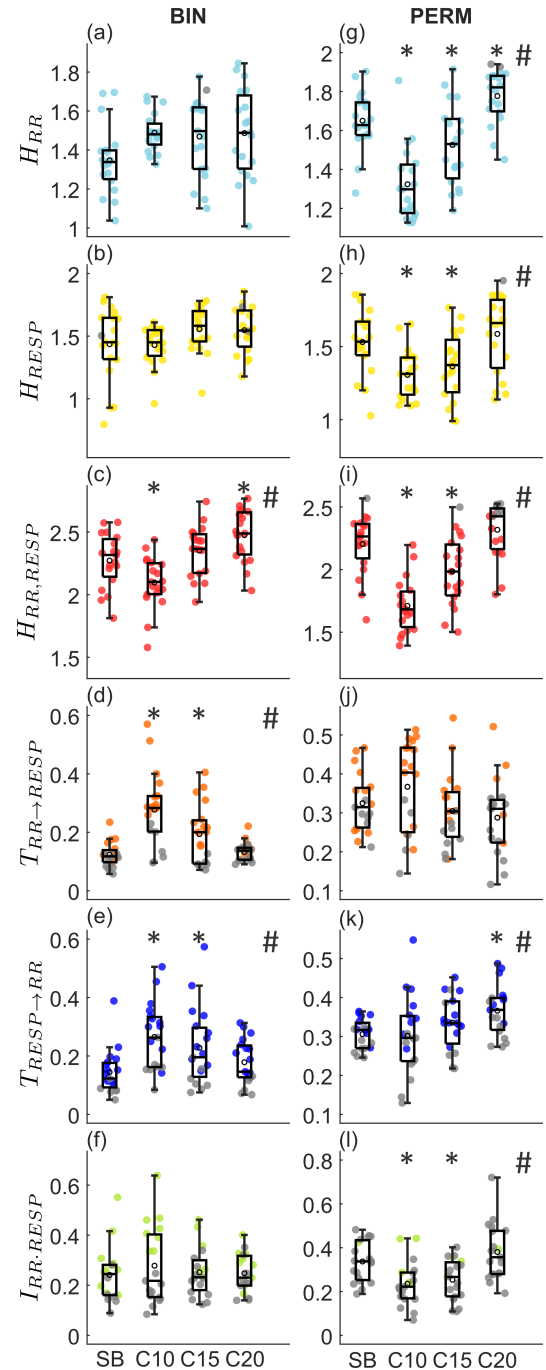


FIG. 4. MIR decomposition of cardiorespiratory interactions (*database I*). Panels report the boxplot distributions and individual values of the MIR decomposition evidencing conditional entropy terms (Eq. (3), H_{RR} , H_{RESP} , $H_{RR,RESP}$) or the MIR decomposition evidencing conditional mutual information terms (Eq. (4), $T_{RR \rightarrow RESP}$, $T_{RESP \rightarrow RR}$, $I_{RR-RESP}$), computed during spontaneous breathing (SB) and controlled breathing at 10, 15, 20 breaths/minute (C10, C15, C20) using the binning approach (a-f) or the permutation approach (g-l) for discretization. In each panel, colored and gray circles correspond to values deemed respectively as statistically significant and non-significant using surrogate data analysis; black open circles refer to the sample mean of each distribution. Statistically significant differences: #, $p < 0.05$, ANOVA; *, $p < 0.05/3$ SB vs. C10/C15/C20, paired t-test with Bonferroni correction.

TABLE I. Results of statistical analyses (ANOVA and uncorrected t-test p -values) and effect size (Cohen's d) performed on cardiorespiratory variability series (*database 1*). Values in bold indicate statistically significant ANOVA ($p < 0.05$), post-hoc t-test p -values after Bonferroni correction ($n = 3$) and large ($|d| > 0.8$) effect sizes. t-test p -values are not in bold if the ANOVA test is not significant.

Measure	Type	p ANOVA	p SB-C10	d SB-C10	p SB-C15	d SB-C15	p SB-C20	d SB-C20
H_{RR}	binning	0.0568	0.0008	-1.005	0.0114	-0.645	0.0141	-0.688
	permutation	$<10^{-9}$	$<10^{-7}$	1.908	0.0115	0.683	0.0046	-0.878
H_{RESP}	binning	0.1136	0.9331	0.028	0.0901	-0.496	0.1724	-0.458
	permutation	0.0005	0.0012	1.136	0.0110	0.731	0.3412	-0.218
$H_{RR,RESP}$	binning	$<10^{-5}$	0.0051	0.830	0.1890	-0.383	0.0058	-0.970
	permutation	$<10^{-10}$	$<10^{-6}$	2.186	0.0019	0.853	0.0874	-0.498
$T_{RR \rightarrow RESP}$	binning	$<10^{-6}$	0.0003	-1.617	0.0126	-0.901	0.5162	-0.235
	permutation	0.0838	0.1689	-0.424	0.4190	0.235	0.1667	0.421
$T_{RESP \rightarrow RR}$	binning	0.0028	0.0041	-1.245	0.0080	-0.780	0.1055	-0.443
	permutation	0.0260	0.8766	0.051	0.0788	-0.529	0.0015	-1.128
$I_{RR,RESP}$	binning	0.7551	0.2294	-0.279	0.6785	-0.102	0.7901	-0.071
	permutation	0.0001	0.0137	1.049	0.0138	0.895	0.2093	-0.387
$I_{RR;RESP}$	binning	$<10^{-5}$	0.0002	-1.573	0.0056	-0.843	0.2540	-0.305
	permutation	0.0141	0.0952	0.500	0.1111	0.502	0.1692	-0.412

II). Physiologically, these results find large confirmation in the literature and are related to the tilt-induced activation of the sympathetic nervous system which tends to simplify the cardiac dynamics, yielding a reduction of the complexity of short-term heart rate variability^{7,18,23,57}.

As regards the causality measures computed via conditional mutual information, the most straightforward result is the increase of the transfer entropy from SAP to RR moving from rest to tilt, which is documented by both binning and permutation approaches (Fig. 5e,k). The information transfer along the opposite direction from RR to SAP shows a significant decrease during tilt, which is detected using the permutation method (Fig. 5j) but not using the binning method (Fig. 5d). According to both discretization approaches, the information shared instantaneously by the two series is rather low and significant in a few subjects, and tends to decrease significantly while moving from rest to tilt (Fig. 5f,l). These results show how the information transferred within the RR-SAP closed-loop is balanced along the two directions of interaction in the resting supine position, and becomes unbalanced in the upright position as a consequence of the well-known tilt-induced activation of the baroreflex feedback from SAP to RR accompanied by a weakening of the mechanical feedforward from RR to SAP^{6,7,46,58}.

4. MIR decomposition: cerebrovascular variability analysis

Fig. 6 depicts the distribution of the information measures composing the MIR computed for the MAP and MCBFV time series measured during the rest and tilt conditions using the binning (a-f) and permutation-based (g-l) approaches.

The MIR decomposition based on conditional entropy measures performed with the permutation approach evidences a statistically significant decrease of the entropy rate of MAP variability (Fig. 6g) and of the joint entropy rate of MAP and MCBFV (Fig. 6i) moving from rest to tilt, while the individ-

TABLE II. Results of statistical analyses and effect size (Cohen's d) performed on cardiovascular variability series (*database 2*). Values in bold indicate statistically significant p -values and large ($|d| > 0.8$) effect sizes.

Measure	Type	p R-T	d R-T
H_{RR}	binning	$<10^{-4}$	2.103
	permutation	0.0053	1.166
H_{SAP}	binning	0.9214	-0.030
	permutation	0.0554	0.704
$H_{RR,SAP}$	binning	0.0006	1.563
	permutation	0.0012	1.370
$T_{RR \rightarrow SAP}$	binning	0.9916	-0.003
	permutation	0.0267	0.766
$T_{SAP \rightarrow RR}$	binning	0.0002	-1.515
	permutation	0.0260	-1.124
$I_{RR,SAP}$	binning	0.0289	0.782
	permutation	0.0027	1.332
$I_{RR;SAP}$	binning	0.1895	-0.366
	permutation	0.1030	0.705

ual entropy rate of MCBFV does not change between the two conditions (Fig. 6h). This response to postural stress, which could not be detected using the binning estimator (Figs. 6a-c) is indicative of an increase of the predictability of the cerebrovascular dynamics, which is driven by a larger predictability of MAP but not MCBFV. As regards the MIR decomposition based on conditional mutual information, all its terms estimated either via the binning or via the permutation approach are not statistically significant in several subjects and do not change significantly moving from rest to tilt (Figs. 6d-f, 6j-l). This result may indicate the presence of low coupling between MAP and MCBFV, or the limited ability of the considered discretization strategies to detect such coupling in the analyzed dataset.

The results presented above are in agreement with the hypothesis that the physiological mechanisms related to cere-

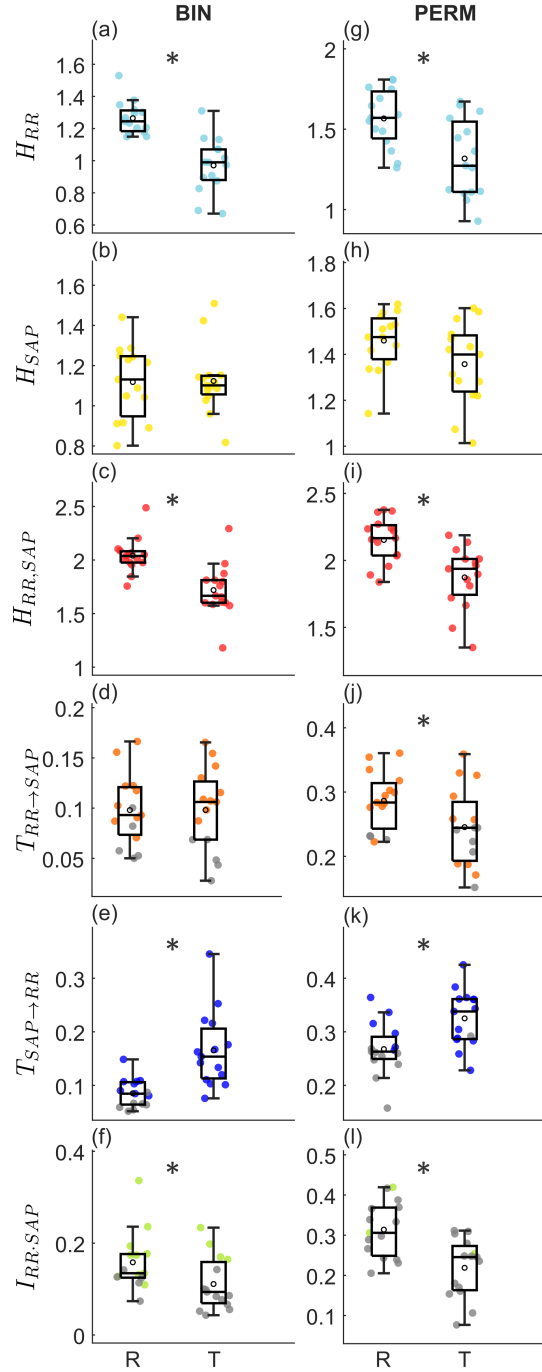


FIG. 5. MIR decomposition of cardiovascular interactions (*database 2*). Panels report the boxplot distributions and individual values of the MIR decomposition evidencing conditional entropy terms (Eq. (3), H_{RR} , H_{SAP} , $H_{RR,SAP}$) or the MIR decomposition evidencing conditional mutual information terms (Eq. (4), $T_{RR \rightarrow SAP}$, $T_{SAP \rightarrow RR}$, $I_{RR,SAP}$), computed in the resting supine position (R) and in the upright position during tilt (T) using the binning approach (a-f) or the permutation approach (g-l) for discretization. In each panel, colored and gray circles correspond to values deemed respectively as statistically significant and non-significant using surrogate data analysis; black open circles refer to the sample mean of each distribution. Statistically significant differences: *, $p < 0.05$ R vs. T, paired t-test.

TABLE III. Results of statistical analyses and effect size (Cohen's d) performed on cerebrovascular variability series (*database 3*). Values in bold indicate statistically significant p -values and large ($|d| > 0.8$) effect sizes.

Measure	Type	p R-T	d R-T
H_{MAP}	binning	0.4028	-0.260
	permutation	0.0015	1.404
H_{MCBFV}	binning	0.3992	0.346
	permutation	0.3738	0.261
$H_{MAP,MVBFV}$	binning	0.5355	0.235
	permutation	0.0020	1.060
$T_{MAP \rightarrow MCBFV}$	binning	0.3854	0.334
	permutation	0.1701	0.597
$T_{MCBFV \rightarrow MAP}$	binning	0.3186	-0.381
	permutation	0.2836	0.434
$I_{MAP-MCBFV}$	binning	0.2608	-0.366
	permutation	0.9519	0.020
$I_{MAP:MCBFV}$	binning	0.4185	-0.300
	permutation	0.2956	0.462

brovascular autoregulation, which are responsible for regulating the cerebral blood flow and maintaining it almost constant independently from changes in the systemic arterial pressure^{59,60}, are preserved after postural stress. Similar findings were obtained by recent studies suggesting that the increased sympathetic nerve activity occurring with the postural challenge leads to a reduction of MCBFV, but not to changes in its variability or predictability, and does not alter significantly the coupling between MCBFV and MAP^{24,25,34}.

IV. DISCUSSION

The purpose of this work was to assess the effectiveness of model-free symbolization methods for assessing coupled physiological dynamics through the MIR measure and its decomposition terms. A comparative investigation of symbolic methods based on binning and permutation was performed to evaluate complex dynamics and dynamical interactions in short-term cardiorespiratory, cardiovascular, and cerebrovascular time series under different physiological states, and the results achieved delineate advantages and drawbacks of each method.

A. Estimating the Decomposition of the Mutual Information Rate via model-free discretization methods

While the theoretical formulation of the MIR (Eq. (1)) and of its decompositions evidencing conditional entropy measures (Eq. (3)) and conditional mutual information measures (Eq. (4)) are well-defined and long-known^{26,28-30}, the practical estimation of all these measures from finite-length realizations of coupled random processes is far from trivial. A main issue is related to the fact that both the MIR and the terms of its decomposition are defined for infinite-dimensional variables sampling the whole temporal evolution of the analyzed pro-

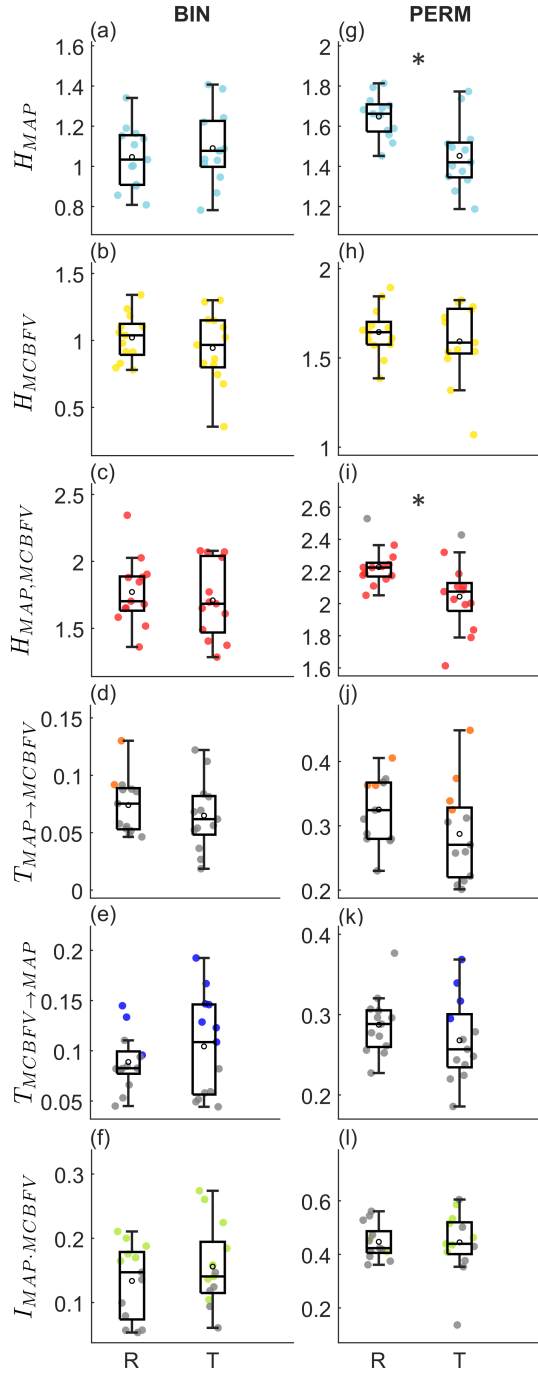


FIG. 6. MIR decomposition of cerebrovascular interactions (*database 3*). Panels report the boxplot distributions and individual values of the MIR decomposition evidencing conditional entropy terms (Eq. (3), H_{MAP} , H_{MCBFV} , $H_{MAP, MCBFV}$) or the MIR decomposition evidencing conditional mutual information terms (Eq. (4), $T_{MAP \rightarrow MCBFV}$, $T_{MCBFV \rightarrow MAP}$, $I_{MAP, MCBFV}$), computed in the resting supine position (R) and in the upright position during tilt (T) using the binning approach (a-f) or the permutation approach (g-l) for discretization. In each panel, colored and gray circles correspond to values deemed respectively as statistically significant and non-significant using surrogate data analysis; black open circles refer to the sample mean of each distribution. Statistically significant differences: *, $p < 0.05$ R vs. T, paired t-test.

cesses (see Eqs. (1,2,5)). This issue is often addressed using linear parametric models to describe the interactions between the present and past states of the processes. In such a case, under the assumption of Gaussian multivariate process the information dynamical measures are directly derived from the covariance matrices of the processes. Thus, the conditional entropy of the present of a process given its past and/or the past of the other process is defined in terms of the variance of the prediction error of a linear regression of the present variable on the past variables. The relations connecting e.g. the entropy rate to the prediction error variance³⁵ and the transfer entropy to Granger causality⁶¹, can be practically computed even on short time series accounting also for long memory effects by employing parametric approaches which involve the solution of the Yule-Walker equations^{55,62} or the formulation of state space models^{32,63}, or even exploiting the relations between time-domain and spectral measures^{28,34,64}.

However, when a model-free analysis is required to account for possible nonlinear interactions, which invalidate the Gaussian assumption, the only viable approach is to work under the Markov property, reducing drastically the length of the analyzed dynamical patterns by setting finite values (at an order depending on the time series length) for the parameter related to memory effects (k in Eqs. (1,2) and m in the approximation of the measures in Eq. (5)) and assuming that longer-memory effects are negligible. In model-free approaches like those based on discretization used in this work, the use of short memories is necessary to limit the curse of dimensionality⁴⁷, i.e., the inability to make sense of a small set of points in a high dimensional space. In particular, estimation of probability distributions computed for high-dimensional variables observed from short time series is problematic leading to highly biased entropy estimates and entropy rate estimates that decrease toward zero as the memory of the process increases^{33,37}. Moreover, an additional bias results from the combination of entropy terms computed on variables spanning spaces of different dimensions^{65,66}; this issue is exacerbated in the estimation of the MIR, as implemented here, which results combining several entropy terms estimated independently from one another. In contrast to the binning and permutation entropy estimates, the nearest neighbor estimate corrects for this bias by taking the spaces regarding the different variables in the entropy terms as projections of the largest space^{65,67}.

In this work, the above issues are reflected by the necessity to choose values of the estimator parameters which limit the dimension of the spaces populated by the points formed by present and past values of the analyzed processes. Specifically, given the challenging conditions of short-term variability where the time series length was in the range of 250-300 points, the memory length was set to $m = 3$ for the permutation approach and to just $m = 1$ for the binning approach; for the latter estimator, a number of quantization levels $b = 4$ was also chosen to limit the alphabet size in the worst condition. While the results obtained for the various MIR terms are physiologically plausible in all the three application contexts considered, some of our findings suggest caution in the use of these parameters or to seek for alternative ways to set

them (e.g., non-uniform embedding^{67–69}). In fact, the lack of statistical significance of $T_{X \rightarrow Y}$, $T_{Y \rightarrow X}$, $I_{X:Y}$ and $I_{Y:X}$ observed in many subjects in all applications can be ascribed to the combination of biases obtained working with high dimensional spaces and small data size. The chosen parameters make this behavior most visible for the permutation-based estimates of instantaneous causality and MIR measures since this is the highest-dimensional case for which the empirical rule suggested to limit the bias is not respected. In the case of binning, the necessity of setting low values of m and b in order to keep the alphabet sizes low and comparable for the different entropy measures appearing in (5) is likely the reason for some unexpected behaviors of the MIR terms, as we will discuss in the next section.

B. Comparison of binning and permutation-based approaches

The binning method is perhaps the most intuitive method to compute entropy measures via discretization, and has been used in numerous previous works analyzing short-term physiological interactions^{20,33,37,68}. As mentioned before, in order to mitigate the issue related to the number of symbolic patterns used to describe the observed dynamics, a low number of quantization levels ($b = 4$) and a limited embedding dimension ($m = 1$) have been set to deal with the small size of the analyzed series (250–300 points). These choices have the consequence to limit the degree of detail in analyzing the amplitude variations and the time-lagged interactions occurring within and between the analyzed processes, possibly impairing the ability to detect subtle changes and long-memory effects. On the other hand, the permutation method allows to work with a number of symbolized patterns which is typically much lower than that used for binning^{43,44}. This occurs because, given the same embedding dimension, the rank-ordering procedure implicitly leads to a much lower number of patterns than the quantization procedure. Considering this aspect, which is expected to improve the estimation of probability distributions and thus of entropy measures, a higher embedding dimension ($m = 3$) has been set when applying the permutation-based method, which allows to achieve a more detailed description of the memory effects. As described in the following, in our analyses the selection of the discretization parameters was reflected on the effectiveness of the estimated information measures regarding the detection of physiological changes, to an extent depending on the application considered.

In cardiorespiratory variability analysis, the permutation-based estimates were found to better reflect the modifications of physiological interactions during the phases of the breathing protocol. Indeed, the fact that RSA is more evident during paced breathing at low breathing frequency⁵² is reflected by an increase of the predictability and a decrease of the complexity of heart rate variability compared to spontaneous breathing, and this effect becomes less pronounced as the breathing rate increases⁵⁵. Conversely, no evident changes in the information transferred between heart rate and respiration variability have been detected using either linear and nonlinear measures of Granger causality^{55,56}. These patterns

were detected in our analysis using permutations but not using the binning approach. This difference may be ascribed to the utilization of a too short embedding dimension, which prevents the binning approach to detect the decreased complexity of RR and RESP during slow paced breathing, possibly because the memory effects occur at lags $m > 1$. A confirmation of this comes from the fact that, when measures working in higher dimensions such as the joint entropy rate or the transfer entropies were computed, significant variations during slow paced breathing were detected also by the binning estimator (Fig. 4); in the case of the transfer entropies the significant variation is likely misleading and reflects more the use of a larger dimension than a physiological behavior. These considerations were supported repeating the analysis of cardiorespiratory interactions based on binning and observing patterns of information dynamics more adherent to those already found in the literature^{55,56} and in this study using the permutation method (results not shown).

As regards the analysis of cardiovascular interactions, we found that the binning approach leads to results similar to those achieved by the permutation method and by other linear and nonlinear measures of complexity, coupling and causality^{16,46,68,70}. Here, our results evidenced that both permutation and binning methods are able to reflect marked variations in the cardiovascular dynamics occurring during postural stress, such as the drop in the complexity of RR and the rise of the causal coupling from SAP to RR (Fig. 5). The only physiological effect that was not detected by the binning method is the decrease of $T_{RR \rightarrow SAP}$ during tilt. Again, this could be ascribed to the limited memory used to evaluate the past dynamics of the process. In fact, lag-specific approaches employed on the same dataset to assess causal RR-SAP interactions suggested that the transfer of information occurs primarily at lags 0 and 1 in the direction from SAP to RR and at lag 2 in the opposite direction from RR to SAP^{20,68}.

The observations above point out that the permutation approach yields results that are more in line with the literature than those provided by the binning approach, mainly as result of the fact that the more parsimonious representation of the discretized dynamic patterns allows to explore higher-dimensional spaces. Nevertheless, the use of $m = 3$ yields a number of patterns which is lower than the time series length when the individual entropy rates or transfer entropies are estimated, but approximately double the series length when the measures involving the highest dimensional spaces are involved. This size for the alphabets turned out to be problematic particularly for the instantaneous information shared by the processes, which resulted barely significant in all applications when computed using permutations. Moreover, the permutation-based method suffers from some limitations that may influence its ability to capture accurately the analyzed dynamic interactions. First, this method is affected by a bias due to the presence of equal values inside the patterns which can confound the estimation⁷¹. While in our work we decided that equal values are ranked according to their temporal order, other choices can be made but none of them actually limits the bias in the estimation of different quantities. With regard to this, a modified permutation-entropy has been proposed which

maps equal values to the same symbols⁷², so as to take into account the presence of equal values within the patterns that may have physiological significance, but this leads to an increase of the dimension of the space visited by the discrete variables and thus to the above-discussed matters. The second constraint of the permutation method is the lack of information related to differences in the amplitude among the time series samples, which may represent a limitation in following some physiological nonlinear behaviors⁷³, but also in the presence of noise⁷⁰. In fact, it was found that permutation-based measures are more susceptible to broad-band noise especially when working with low signal-to-noise ratio and short time series length. On the other hand, not accounting for signal amplitudes makes the permutation method more robust in the presence of nonstationarities^{38,44}. In particular, the presence of amplitude artifacts and slow trends, which is highly detrimental for most entropy estimators⁴⁰, should have a lower impact on procedures which discard the amplitude values apart from considering their rank. This aspect may be relevant in our application to cerebrovascular dynamics, where the occurrence of physiological oscillations at low and very low frequency trends³⁴ can be one of the reasons for the inability of the binning method to capture the complexity changes of mean arterial pressure evoked by tilt.

C. Conclusion and future investigations

This study points out the feasibility of investigating short-term interactions in bivariate physiological time series by means of model-free estimators of the mutual information rate and of its constituent terms implemented via discretization strategies. The most critical aspect identified by our results is the need to set the analysis parameters to values that limit the curse of dimensionality. From this point of view, the permutation approach to discretization based on rank ordering is preferable, as it yields a more parsimonious representation of the discretized patterns representing the observed dynamics; still, the use of permutations becomes problematic when implemented in high-dimensional spaces. As regards the binning approach, we suggest that it can be used with the standard setting of the embedding and quantization parameters for estimating the entropy rate of individual time series, while shorter memories and heavier coarse graining should be used to reliably assess the higher-dimensional terms of the MIR decomposition.

Further work is envisaged to explore more in depth the performance of model-free approaches for the estimation of dynamic information measures from short realizations of multivariate time series, and to develop improved estimators. In particular, future studies should assess more systematically the behaviours of binning and permutation approaches on simulated dynamic systems, also in comparison with approaches not based on discretization developed using either model-free^{65,66} or model-based^{7,11} estimators. In parallel, strategies for dimension reduction^{66,67,69} or automatic parameter selection⁷⁴ should be explored to optimize the embedding of multivariate time series limiting the curse of dimensionality.

In perspective, the combination of improved entropy estimators and optimized methods for parameter selection can open new avenues for the model-free assessment of the information processed by network systems, even beyond the framework of pairwise interactions^{32,75}.

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DATA AVAILABILITY STATEMENT

The data presented in this study will be made available upon reasonable request from the corresponding author.

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