

1 Prediction of debris-avalanches and -flows triggered by a
2 tropical storm by using a stochastic approach: An
3 application to the events occurred in Mocoa (Colombia) on
4 1 April 2017

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14 A B S T R A C T

15 Landslides are among the most dangerous natural processes. Debris avalanches and debris flows in
16 particular have often caused casualties and severe damage to infrastructures in a wide range of
17 environments. The assessment of susceptibility to these phenomena may help policy makers in
18 mitigating the associated risk and thus it has attracted special attention in the last decades.

19 In this experiment, we assessed susceptibility to debris-avalanche and -flow landslides by using a
20 stochastic approach. Two different modeling techniques were employed: i) Multivariate Adaptive
21 Regression Splines (MARS) and ii) Logistic Regression (LR). Both MARS and LR allow for
22 calculating the probability of landslide occurrence by building statistical relationships between a set
23 of environmental variables and the target variable, i.e. presence/absence of the landslide event. The

24 target variable was extracted from an inventory of debris-avalanche and -flow landslides which
25 were triggered by the tropical storm that hit the area of Mocoa (Colombia) on 1 April 2017. As
26 predictor variables, we employed nine terrain attributes derived from a 5-m resolution DEM (i.e.
27 elevation, slope angle, northness, eastness, upslope slope angle, convergence index, topographic
28 position index, valley depth and topographic wetness index), in addition to lithology, distance from
29 faults and presence/absence of soil creep processes. In our experiment, we used three different
30 landslide datasets which contain i) the highest point of each recognized landslide crown-lines
31 (dataset LIP), ii) the highest 10% of cells of each landslide area (dataset SOURCE), and iii) the
32 entire landslide areas, which include initiation and accumulation zones (dataset MASS). In order to
33 evaluate their predictive ability, LR and MARS models were submitted to k -fold spatial cross-
34 validation strategy, which consists in extracting random training and test subsets from k spatially
35 disjoint sub-areas. The results of model validation, expressed in terms of Area Under the ROC
36 Curve (AUC), demonstrate better predictive performance of MARS models with respect to LR
37 models, for all the three landslide datasets. The mean AUC values calculated for the datasets LIP,
38 SOURCE and MASS of the MARS models are 0.776, 0.788 and 0.768, respectively, whereas AUC
39 values of the LR models are 0.748, 0.751 and 0.703, respectively. Models validation also show that
40 the predictive skill of the models is better when landslide data are sampled from the highest
41 portions of the landslides (dataset SOURCE). Maps of susceptibility to debris-avalanche and -flow
42 landslides for the Mocoa area were produced by using both LR and MARS and the three landslide
43 datasets. The analysis of the distribution of events versus the susceptibility classes of the maps
44 confirm that MARS and the dataset SOURCE provide the best ability to discriminate between event
45 and non-event cells.

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47 *Keywords:* Debris flows; Landslide susceptibility; Tropical storm; Multivariate Adaptive
48 Regression Splines (MARS); Logistic Regression (LR); Mocoa (Colombia).

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52 **1. Introduction**

53 In the last decades, the population growth and the urbanization of hazardous areas have largely
54 increased the damage and loss of lives due to natural disasters. In many developed and developing
55 countries, landslides are among the most important causes of natural hazard (Guzzetti et al., 1999).
56 Landslide hazard is particularly important in mountainous environments where, due to topography,
57 these phenomena may achieve rapid propagation and high energy. Some of the most devastating
58 landslide disasters are related to the occurrence of debris avalanches and debris flows triggered by
59 heavy rainstorms or earthquakes (Hungri et al., 2014). When caused by intense and prolonged
60 rainfalls, these landslides may occur simultaneously with hyperconcentrated flows and flash floods.
61 A dramatic example of their destructive potential is given by the disaster of Vargas (Venezuela)
62 which caused around 15,000 fatalities in December 1999 on a narrow coastal zone north of Caracas
63 (Larsen and Wicczorek, 2006; Larsen, 2008). More recently, in January 2011, over 1,500 casualties
64 were caused by the disaster occurred in the mountainous region of Rio de Janeiro (Brazil), where
65 clusters of debris avalanches and debris flows were triggered by an extreme rainfall event during a
66 period of 2 days (Avelar et al., 2013, Hungri et al., 2014). Other examples of natural disasters that
67 occurred worldwide and related to debris avalanche/flow landslides are reported in the literature
68 (e.g., Crosta and Dal Negro, 2003; Crozier, 2005; Aronica et al., 2012).

69 Mitigation of landslide risk may be achieved by predicting where landslides are more likely to
70 occur in the future. This information could indeed help policy-makers in implementing land-use
71 strategies aimed at minimizing human casualties and property damage (Guzzetti et al., 1999). The
72 likelihood of landslide occurrence in a given area is defined as landslide susceptibility (Brabb,
73 1984; Carrara et al., 1995). A landslide susceptibility map depicts the spatial relative probability of
74 landslide occurrence within a given area (Conoscenti et al., 2016).

75 Landslide susceptibility mapping can be achieved by using different methods. Among them, the
76 stochastic approach has become very popular over the last decades, due to the availability of high-
77 resolution terrain data and freeware statistical and Geographical Information Systems (GIS)

78 software. This approach is based on the assumption that “the past is the key to the future” (Carrara
79 et al., 1995) and new slope failures are more likely to occur under the same conditions that caused
80 landsliding in the past. Therefore, statistical methods require the location of past landslides and
81 maps of environmental variables which directly or indirectly (as proxies) reflect the preparatory
82 factors controlling landslide occurrence in the study area. Different statistical and data mining
83 modeling techniques have been proved to provide reliable and accurate landslide susceptibility
84 maps starting from event inventories and sets of predictor variables (Aleotti and Chowdhury, 1999;
85 Guzzetti et al., 1999; Brenning, 2005; Reichenbach et al., 2018).

86 Although widely adopted in recent years, statistical modeling of landslide susceptibility involves
87 some critical issues which still remain unsolved. One of these is related to the selection of event and
88 non-event data which are employed to calibrate and validate the landslide predictive models. This is
89 recognized as a crucial step influencing the accuracy and reliability of the final landslide
90 susceptibility models and maps.

91 In case of grid-based landslide predictions, non-event locations are typically sampled from stable
92 portions of slopes, i.e. cells outside the landslide areas. On the other hand, no agreement exists on
93 the best approach to select landslide cells. These should identify landscape locations where levels of
94 the landslide controlling factors exceeded the threshold of slope stability and, thus, triggered the
95 slope failure. However, identifying the exact initiation points of a large number of landslides is very
96 problematic (Regmi et al., 2014). In the case of debris avalanches or flows, even distinguishing
97 between source and accumulation zones could be quite difficult. Furthermore, once initiated, a
98 landslide can extend upslope, downslope and/or sidewise, making the identification of the initiation
99 point a very challenging task.

100 Various landslide data selection and sampling techniques have been proposed to minimize the
101 uncertainty in identifying landslide initiation points. The most frequently adopted strategies are: i)
102 single cells randomly selected from landslide areas or from depletion zones (e.g., Vorpahl et al.,
103 2012; Heckmann et al., 2014; Goetz et al., 2015); ii) centroid of landslide areas or of depletion

104 zones (e.g., Atkinson and Massari, 2011; Regmi et al., 2014); iii) multiple cells (all or a fraction)
105 within landslide areas or depletion zones (e.g., Regmi et al., 2014; Conoscenti et al., 2016); iv)
106 single cell or all the cells in the upper edge of the main scarp (e.g., Clerici et al., 2006; Costanzo et
107 al., 2012; Cama et al., 2015; 2017). Moreover, when the available DEM is more recent than the
108 landslide inventory, topographic triggering conditions are sampled from a buffer around the
109 landslide polygons (e.g., Süzen and Doyuran, 2004; Nefeslioglu et al. 2008; Rotigliano et al., 2011)
110 or from a reconstructed pre-failure topography within landslide areas (e.g., Van Den Eeckhaut et al.
111 2006; Gorum et al. 2008; Conoscenti et al., 2015).

112 Intense and prolonged rainfall events occur episodically in tropical Andes, triggering debris-
113 avalanche and -flow landslides, hyperconcentrated flows and flash floods. These natural hazards
114 threaten many communities living on or near alluvial fans where they may cause fatal victims and
115 extensive property damage. The assessment of susceptibility to debris-avalanche and -flow
116 landslides in these environments is therefore crucial to plan risk mitigation strategies and prevent
117 large disasters.

118 On 1 April 2017, the area of Mocoa (Colombia) was hit by a sever tropical storm which discharged
119 130 mm of rain in 3 h starting from 10:00 pm of March, 31st. This heavy rainfall triggered more
120 than one thousand debris-avalanche and -flow landslides, which evolved to hyperconcentrated
121 flows in the main streams of the Mulato, Sangoyaco and Taruca rivers, three tributaries of the
122 Mocoa river. The event caused 328 victims, 200 missing, more than 1000 injured and damage to
123 more than 120 houses in 17 neighborhoods of the city.

124 In this study we focused on the prediction of the rainfall-induced debris-avalanche and -flow
125 landslides which occurred in the area of Mocoa. These processes initiated on the slopes and
126 provided sediments to hyperconcentrated flows which moved rapidly down steep channels and
127 caused debris flooding of large parts of the city located on the alluvial fans. Within the context of
128 testing different methods of landslide data sampling for susceptibility modeling, we used and
129 compared three criteria which differ in the way the event data were selected. The landslide datasets

130 employed in our experiment include (i) the highest point of each landslide area, (ii) the highest 10%
131 of cells of the landslide areas or (iii) the entire landslide areas. To predict the spatial distribution of
132 the Mocoa landslides, we used two different statistical modeling techniques, namely: (i) logistic
133 regression (LR) and (ii) multivariate adaptive regression splines (MARS). LR has been widely
134 adopted to assess landslide susceptibility and, in particular, debris avalanche/flow landslides.
135 Conversely, MARS has been rarely used in the field of landslide susceptibility mapping and, as far
136 as we know, has been employed to predict these types of slope failures only in one research study
137 (Rotigliano et al., 2018).

138 Therefore, the main objectives of this experiment were to: (i) test and compare three different
139 criteria of landslide data sampling; (ii) evaluate and compare the ability of LR and MARS to predict
140 the debris-avalanche and -flow landslides occurred in Mocoa.

141 **2. Study area**

142 The study area is located in the southeast portion of the Eastern Cordillera of the Colombian Andes
143 and falls in the watershed of the Mocoa River (Fig. 1). It extends for 58.3 km² and lies between
144 latitudes 01° 08' and 01° 13' N and longitudes 76° 38' and 76° 42' W. The city of Mocoa, capital
145 of the Departamento de Putumayo, is located in the southeast sector of the study area. Its urban area
146 is bounded to the east by the Mocoa river and is crossed by three of its tributaries: Taruca,
147 Sangoyaco and Mulato. Precipitation in the study area occurs all year long and shows a unimodal
148 annual pattern with highest and lowest average monthly rainfall occurring in June and January,
149 respectively. The average annual rainfall is approximately 3715 mm.

150 The topography of the study area varies from flat to hilly to steep, with elevation ranging from 538
151 to 1893 m asl (Figs. 2 and 3). The western portion is characterized by moderately dissected
152 mountain chains with very steep slopes (up to 77°). These are followed to the east by a hilly sector
153 where topography is mainly controlled by structure and the slope angle is in the range 10–25°.
154 Gentle-sloping (<10°) surfaces, which are formed by coalescing debris fans, occur on the right
155 Mocoa riverside. A structure-controlled relief, which is located to the north of the Mocoa urban

156 area, acts as a natural barrier to protect the city from flooding events. Moreover, flat surface (i.e.
157 river terraces and alluvial plains) are located along the inter-mountainous sector of the Mocoa river.
158 The study area is underlain by Jurassic igneous rocks and Cretaceous and Paleocene-Eocene
159 sedimentary rocks, which are covered discordantly by Quaternary unconsolidated deposits. To
160 explore the role of lithology as a predictor of landslide distribution, we prepared a map that includes
161 the following lithological units: i) sedimentary rocks, mainly conglomerates; ii) sedimentary rocks,
162 mainly calcareous limestones; iii) igneous rocks mainly granites and monzonites; iv) alluvial
163 deposits; v) terraced alluvial deposits; vi) colluvial deposits; vii) debris-torrent deposits (Fig. 4).
164 The igneous rocks, which are moderately to highly fractured, crop out in the western sector of the
165 study area. The La Tebaida fault separates these rocks to the east from the Paleocene-Eocene
166 conglomerates of the Pepino Fm., whereas the limestones of the Villeta Fm. outcrop in the eastern
167 side of the Mocoa River. The quaternary deposits occupy the Mocoa river bed and form fluvial
168 terraces along its main valley. Moreover debris fans occur along the Mocoa tributaries while
169 colluvial deposits cover the foot of some slopes along the inter-mountain valleys.

170 **3. Materials and methods**

171 *3.1. Landslide inventory*

172 The rainfall event considered in this study occurred between 10 pm of March, 31st and 1 am of
173 April, 1st, 2017. The event was short but very intense with 130 mm of rain that fell in about 3 h.
174 Such an amount of rain usually occurs in a ten days time-lapse. The rainfall event triggered a high
175 number of debris avalanches and debris flows which propagated downslope as hyperconcentrated
176 flows and debris floods.

177 The landslides occurred in the slopes that fed the streams of the Taruca, Sangoyaco and Mulatos
178 rivers with a large amount of debris and blocks (with diameter up to six meters). These sediments
179 were transported downslope and covered an area of approximately 3.2 km², with average thickness
180 of 4 m and a total volume of around 12 millions of cubic meters.

181 In this work, we built a landslide inventory by 2D and 3D visual analysis of aerial photographs and
182 of the satellite image dated April 10th, 2017, available on Google Earth. The inventory includes
183 1347 landslides (Fig. 5). Based on Hungr et al. (2014), two main types of movements were
184 recognized: i) debris avalanches, which occurred on steep slopes and without confinement in an
185 established channel and ii) debris flows, in which the movement developed along established paths,
186 usually first or second order streams.

187 The mapped slope failures probably initiated as shallow slides or flows (Cruden and Varnes, 1996)
188 and after moving a short distance transformed into debris avalanches or debris flows. However,
189 considering that the objective of this experiment was to identify where slope failures potentially
190 reaching the drainage axes and, eventually, the urbanized area of Mocoa, were more likely to occur,
191 we decided not to differentiate between the two types of movements in the susceptibility mapping.

192 3.2. Predictor variables

193 In this experiment, selection of landslide predictors was performed according to quality and
194 resolution of the available data. As spatial distribution of the rainfall event was not available, only
195 variables representing landslide preparatory causes were employed.

196 The landslide predictors were derived from a geological map of the area and a 5-m resolution raster
197 Digital Elevation Model (DEM), which provided sufficient resolution to properly map susceptibility
198 to landsliding in the Mocoa area. Conversely, the resolution of existing land cover and soil maps
199 were too coarse and thus not suitable for the analysis. All the variables were prepared as raster GIS
200 layers with 5-m cell size. The following predictors were derived from the DEM using SAGA-GIS
201 software (Conrad et al., 2015): elevation (*ELEV*), slope angle (*SLOPE*), northness (*NORTH*),
202 eastness (*EAST*), upslope slope angle (*UPSLO*), convergence index (*CONV*), topographic position
203 index (*TPI*), valley depth (*VDEP*) and topographic wetness index (*TWI*). These attributes were
204 selected as proxies for conditions and processes related to landslide occurrence (e.g., Wilson and
205 Gallant, 2000; Ohlmacher, 2007; Vorpahl et al., 2012). *ELEV* reflects the values of the available 5-
206 m DEM and was selected because of its expected correlation with rainfall and vegetation spatial

207 distribution. *SLOPE* was calculated according to Zevenbergen and Thorne (1987). *NORTH* and
208 *EAST* were computed by applying cosine and sine transformations of slope aspect, respectively
209 (Brenning and Trombotto, 2006; Conoscenti et al., 2016; Cama et al., 2017). *NORTH*, *EAST*, as
210 well as *ELEV*, may serve as proxies for seasonal wet/dry cycles of soils (Auslander et al., 2003).
211 *UPSLO* reflects average slope angle upstream from each position in the landscape. *CONV* (Koethe
212 and Lehmeier, 1996), which estimates to what extent neighboring cells point to the center cell, was
213 calculated by setting a search radius of 50-m. To reduce detail and remove noise of the *TWI* (Beven
214 and Kirkby, 1979), a gaussian smoothing filter with search radius of 25-m was applied. *CONV* and
215 *TWI* were included to account for runoff convergence/divergence and potential soil saturation,
216 respectively. *TPI* (Guisan et al., 1999) indicates the relative position of each cell and was calculated
217 by using a 100-m search radius. *TPI* and *CONV* values are negative on valley bottoms and positive
218 on ridges. *VDEP* reflects the maximum relative relief measured in cross-sections and thus is a
219 measure of local relief energy (Lóczy et al., 2012).

220 Moreover, we employed lithology (*LITHO*), distance form faults (*FAULTD*) and presence/absence
221 of soil creep (*CREEP*) as predictor variables. *LITHO* was prepared by grouping the geological
222 formations that outcrop in the study area into 7 geological units, according to their expected
223 relationship with slope stability (Fig. 4). *FAULTD* was included to potentially reflect the degree of
224 weakening of the bedrock due to tectonically active regional geological structures (Mathew et al.,
225 2009; Cama et al., 2017). *CREEP* includes 71 areas (extending between 1 and 285 ha) affected by
226 soil creep (Fig. 5), which in the region is favored by deforestation for agriculture and pasture.
227 *CREEP* was prepared by analyzing the same images employed to map the landslides.

228 3.3. Statistical modeling

229 Probability of debris-avalanche and -flow landslide occurrence at each 5-m grid cell in the study
230 area was calculated by employing two statistical modeling techniques: logistic regression (LR;
231 Hosmer and Lemeshow, 2000) and multivariate adaptive regression splines (MARS; Friedman,
232 1991). The statistical analyses were performed using the R software (R Core Team, 2017) with the

233 packages “usdm” (Naimi, 2015), “sperrorest” (Brenning, 2012) and “earth” (Milborrow et al.,
234 2015).

235 LR is among the most frequently used statistical technique for spatial modeling of landslide
236 occurrence (Brenning, 2005) and it has been employed several times to predict specifically storm-
237 triggered landslides (e.g., Chevalier et al., 2013; Heckmann et al., 2014; Lombardo et al., 2014,
238 2015; Cama et al., 2016; Trigila et al., 2015). On the other hand, MARS has been employed only
239 rarely for assessing landslide susceptibility (e.g., Vorpahl et al., 2012; Felicísimo et al., 2013;
240 Conoscenti et al., 2015, 2016; Pourghasemi and Rossi, 2016) and, as far as we know, this modeling
241 technique was exploited for predicting landslides triggered by extreme rainfall events only in one
242 recent research paper (Rotigliano et al., 2018). LR and MARS can use both continuous and
243 categorical independent variables to estimate a response variable in the range 0 to 1, which can be
244 interpreted as probability of an event occurrence. Both LR and MARS consist of an additive
245 combination of terms. Each term of the LR model is given by a linear regression of an independent
246 variable, which is fitted using the maximum likelihood method. In contrast to the assumption of LR
247 that coefficients of the predictors are constant across their ranges, MARS splits the range of the
248 independent variables into pieces, fitting to each of them a linear regression called “basis function”
249 (Vorpahl et al., 2012; Gómez-Gutiérrez et al., 2015; Conoscenti et al., 2018; Garosi et al., 2018).
250 MARS terms consist of a single basis function or a product of two or more of them. To reduce the
251 complexity of the LR models and avoid problems of overfitting, we adopted a bilateral stepwise
252 strategy that selects only the most significant predictors. Accordingly, MARS models were
253 prepared by avoiding terms consisting of more than one basis function.

254 As both LR and MARS require the absence of multicollinearity (i.e. predictors should not be
255 correlated with each other), we calculated the variance inflation factor (*VIF*) to measure the degree
256 of correlation among the selected predictors. The R package “usdm” was employed to this aim.
257 Following the “rule of 10”, according to which a *VIF* > 10 reveals strong multicollinearity

258 (Heckmann et al., 2014; Jebur et al., 2014; Bui et al., 2015), we decided to include in the models all
259 the above cited variables, as their *VIF* values are well below the threshold.

260 *3.4. Landslide data sampling*

261 In our experiment we used three different methods to select landslide data (Fig. 6). The three
262 methods differ in the way event cells were selected. A single approach was instead employed to
263 pick up non-event cells. These were randomly sampled from stable portions of the slopes, i.e. cells
264 outside landslide areas (dataset STABLE). Landslide data samples were always prepared by
265 maintaining a presence-absence ratio of 1:1. This choice was made in order to avoid prevalence in
266 the samples (i.e. different proportion of event and non-event observations), which has been shown
267 to affect the reliability of common accuracy statistics (Beguería, 2006).

268 In the first approach we selected a single cell at the highest point of each of the recognized landslide
269 crown-lines (hereafter named LIP; Lombardo et al., 2015; Cama et al., 2015, 2016), identifying a
270 total of 1347 event cells (dataset LIP). As model performance and robustness can be affected by the
271 size of calibration and validation datasets (Brenning, 2005; Vorpahl et al., 2012), the same number
272 of event cells was selected also using the other two methods. In our second approach, event cells
273 were randomly sampled from the upper portions of the mapped landslides, which should reflect the
274 conditions of the main source areas. These portions include the highest 10% of cells of each
275 landslide area (dataset SOURCE). In the third method, landslide cells were randomly picked up
276 from the entire landslide areas, which include initiation and accumulation zones (dataset MASS). In
277 the second and third approach, we sampled as event cells a relative small fraction of pixels within
278 both upper portions (around 12% of 11,294 cells) and entire landslide areas (around 1.2% of
279 109,149 cells). In this way, we limited the effects of spatial autocorrelation between sampled cells,
280 which should be avoided when performing statistical prediction of landslide occurrence.

281 3.5. Models training and validation strategy

282 In order to evaluate their predictive ability, LR and MARS models were submitted to a k -fold
283 spatial cross-validation strategy, which consists in extracting random training and test subsets from
284 k spatially disjoint sub-areas. These were identified by using the k -means clustering algorithm (Ruß
285 and Brenning, 2010; Goetz et al., 2015). As classic k -fold cross validation, the adopted approach
286 uses $k - 1$ combined subsets at time for calibration and the remaining one for validation. The
287 process is then repeated k times.

288 In our experiment we used $k = 5$ spatial cross-validation folds. To improve the robustness of the
289 landslide predictions and mitigate the rare-events issue (Heckmann et al., 2014; Svoray et al., 2012;
290 Van Den Eeckhaut et al., 2012), multiple learning samples were extracted for each fold. More in
291 detail, our approach included the following steps, which were repeated for each of the five folds: i)
292 sampling of ten balanced subsets from the four (i.e. $k - 1$) calibration sub-areas; ii) training a model
293 on each of the ten calibration subsets; iii) sampling of a balanced subset from the validation sub-
294 area; iv) calculating a probability (P) of landslide occurrence in the validation subset by averaging
295 the scores obtained from the ten model runs; v) estimation of the model performance by averaging
296 the performance evaluated across each of the five folds. To assess the robustness of our approach,
297 this validation process was repeated 100 times for both LR and MARS models.

298 The prediction skill of the models was evaluated with the area under the receiver operating
299 characteristics (ROC) curve (AUC). The ROC curve plots for all possible cut-off values the true
300 positive rate TPR (sensitivity) versus the false positive rate FPR ($1 - \text{specificity}$). AUC values close
301 to 1 indicate perfect discrimination ability between the target variable levels (0 or 1) whereas values
302 close to 0.5 reflect no discrimination ability of the models. Intermediate AUC values were
303 interpreted as acceptable, excellent or outstanding if higher than 0.7, 0.8 and 0.9, respectively
304 (Hosmer and Lemeshow, 2000). The Wilcoxon signed-rank test was applied to detect significant
305 differences in model performance. Differences at p -value < 0.01 were considered significant.

306 3.6. Debris-avalanche and -flow landslide susceptibility maps

307 A total of six debris-avalanche and -flow landslide susceptibility maps were prepared for the area of
308 Mocoa by using LR and MARS with samples extracted from the datasets LIP, SOURCE and
309 MASS. The following procedure was applied for both the modeling techniques (i.e. LR and MARS)
310 and the three datasets (i.e. LIP, SOURCE and MASS): i) sampling of 100 balanced subsets made of
311 2694 cells; ii) calibration of a model for each of the 100 subsets of cells; iii) calculation of
312 probability (P) of landslide occurrence for each cell of the study area by averaging the scores
313 obtained from the 100 model runs. The range of debris avalanche/flow probability (0.00 – 1.00) was
314 classified into four equal interval levels (interval width: 0.25). As both LR and MARS models were
315 prepared using balanced datasets of event and non-event pixels, the score averaged from the 100
316 model runs should be interpreted as relative probability of debris avalanche/flow occurrence (Goetz
317 et al., 2015).

318 4. Results

319 4.1. Validation of the susceptibility models

320 The predictive performance of LR and MARS evaluated for the three datasets is summarized in Fig.
321 7 by using six box plots. Each box plot shows the variability of 100 AUC values which were
322 computed by means of the validation procedure described in Section 3.5. The average AUC values
323 of LR and MARS calculated for the three datasets are all in the range 0.7 – 0.8, demonstrating an
324 acceptable ($AUC > 0.7$) overall accuracy of the debris-avalanche and -flow landslide predictive
325 models in the area of Mocoa. However, significant differences are revealed if we compare the
326 performance of the models by applying the Wilcoxon signed-rank test.

327 The 100 MARS model runs validated on the dataset SOURCE (mean $AUC = 0.788$) demonstrated a
328 significant (p -value $< 2.2e^{-16}$) higher accuracy than those tested on the dataset LIP (mean $AUC =$
329 0.776), which in turns performed better (p -value $= 7.23e^{-07}$) than the MARS-MASS model runs
330 (mean $AUC = 0.768$). On the other hand, no significant (p -value $= 0.1355$) difference of

331 performance was measured between LR predictions of LIP (mean $AUC = 0.748$) and SOURCE
332 (mean $AUC = 0.750$) datasets whereas LR-MASS model runs clearly exhibited the poorest
333 predictive performance (mean $AUC = 0.703$). In all three datasets, MARS runs outperformed LR
334 model repetitions (p -value $< 2.2e^{-16}$).

335 As regards performance variation, which was evaluated by means of the standard deviation (SD),
336 MARS- and LR-LIP models exhibited very similar robustness (SD: 0.0075 vs 0.0076), whereas
337 MARS predictive skill was slightly more stable when validated on the SOURCE (SD: 0.0098 vs
338 0.0117) and MASS datasets (SD: 0.0102 vs 0.0131). For any of the two modeling techniques, AUC
339 values dispersion increases from LIP, through SOURCE to MASS datasets.

340 4.2. MARS and LR susceptibility maps

341 The six debris-avalanche and -flow landslide susceptibility maps prepared with MARS and LR
342 using the three landslide datasets are plotted in Fig. 8. To help compare and evaluate the maps, Fig.
343 9 shows the relative frequency distributions of the datasets ALL (all pixels in the study area),
344 STABLE, LIP, SOURCE and MASS over the four relative probability classes, for each of the six
345 maps.

346 All the maps show smooth prediction patterns without abrupt changes where boundaries of
347 categorical variables (i.e., lithology and soil creep presence/absence) occur.

348 MARS provides for all landslide datasets a smoother distribution of the susceptibility classes over
349 the entire maps (dataset ALL), with a gradually decreasing frequency from the lowest to the highest
350 class of landslide probability (P). Conversely, except for the SOURCE map, LR produces roughly
351 similar frequency of P classes between 0.00 and 0.75 whereas the highest class occurs more rarely.

352 Non-event pixels (dataset STABLE) are better discriminated by the MARS-SOURCE map, where
353 75% of the stable points has a low susceptibility level (i.e. $P < 0.5$) whereas only 67% has a P value
354 below 0.5 in the LR-LIP map.

355 Fig. 9 shows that, if we consider $P = 0.5$ as threshold to discriminate between pixels predicted as
356 stable ($P < 0.5$) and unstable ($P > 0.5$), event cells of the datasets LIP, SOURCE and MASS are

357 predicted with excellent sensitivity (or true positive rate, TPR) by both MARS (TPR range: 0.77 –
358 0.78) and LR (TPR range: 0.76 – 0.78) maps. However, if we focus on the distribution of event
359 pixels with $P > 0.5$, Fig. 9 reveals that, except for LIP maps, they are roughly equally distributed
360 between the two highest levels of susceptibility in the MARS maps, whereas event pixels are less
361 frequent in the highest P class of LR maps.

362 **5. Discussion**

363 The spatial cross-validation revealed that the statistical predictive models provided an acceptable fit
364 to the spatial distribution of the landslides occurred in the study area during the night between
365 March 31 and April 1, 2017. Although MARS and LR models provided apparently comparable
366 ability to discriminate between event and non-event pixels on the three landslide datasets (i.e. LIP,
367 SOURCE, MASS), the Wilcoxon signed-rank test revealed significant differences of performance.
368 MARS models, indeed, showed a better prediction skill in all landslide datasets, with a mean AUC
369 difference of 0.044.

370 The existence of non-linear relationships between selected predictors and slope failures in the study
371 area can explain the better fit of MARS models to the landslide datasets. LR models are indeed
372 based on linear relationships holding over the entire range of the explanatory variables whereas
373 MARS is able to split the range of predictors into pieces and fit to each of them a different linear
374 regression. In this way, MARS produces a smooth response curve which may better reproduce the
375 relationships between predictors and landslide occurrence.

376 The better performance of MARS in our study area is consistent with the findings of few other
377 previous works that compared the ability of LR and MARS in predicting landslide occurrence. A
378 similar difference of mean AUC values between LR and MARS (0.848 against 0.889) was found by
379 Conoscenti et al. (2015) in a catchment of Sicily (Italy) affected by earth-flow landslides. Roughly
380 the same difference of prediction skill was revealed by internal cross validation that was applied to
381 an inventory of shallow translational landslides in the Andes of Southern Ecuador (Vorpahl et al.,
382 2012). A slightly better performance of MARS ($AUC = 0.782$) with respect to LR ($AUC = 0.775$)

383 was found by Pourghasemi and Rossi (2016), who compared the ability of four statistical
384 techniques to predict landslides occurred in the Mazandarn Province (Iran). On the other hand, LR
385 and MARS achieved the same accuracy ($AUC = 0.76$) in predicting the landslide spatial distribution
386 in Guipúzcoa province (Spain; Felicísimo et al., 2013). Also, Vorpahl et al. (2012) found no
387 significant difference of predictive performance when they applied an external cross validation to
388 their data.

389 The results of our experiment and those of the above cited papers suggest that the role of a physical
390 factor in controlling a geomorphological process can be better represented by a series of local
391 functions rather than a single linear regression. In other words, simple predictive models, which are
392 typically employed to explain relationships between factors and geomorphological processes, are
393 less accurate than complex models, such as those provided by MARS.

394 As regards the different landslide datasets employed in this experiment, we found a significant
395 lower accuracy when predictive models were calibrated and validated using event pixels sampled
396 from the entire landslide areas (i.e. dataset MASS). The best fit to landslide data was found when
397 we used MARS with event cells sampled from the highest 10% of cells of each landslide area
398 (dataset SOURCE) whereas no significant difference of performance was detected when we trained
399 and tested LR models with datasets LIP and SOURCE.

400 The overall better discrimination ability of the LIP/SOURCE pixels can be explained considering
401 that their terrain conditions are more specific than those of the MASS cells. The latter indeed may
402 occur on the initiation, transport and deposition zones of debris-avalanches and -flows which can be
403 characterized by heterogeneous conditions (e.g., altitude, slope, convergence index, TPI, TWI or
404 valley depth) within each individual landslide area, whereas LIP/SOURCE pixels should likely
405 have more homogeneous conditions. These results are consistent with those of Regmi et al. (2014),
406 who found that accuracy of landslide predictive models was slightly better when developed on
407 samples obtained from scarps of different types of landslides (i.e., debris flows, debris slides, rock
408 slides and soil slides) occurred in western Colorado, USA. They explain their findings assuming

409 that landslides in their study area were mainly due to unfavorable condition at the slope heads and,
410 thus, data sampling from scarps instead of landslide masses reduced uncertainties. On the other
411 hand, Vorpahl et al. (2012) found better performance in predicting landslide deposition zones than
412 initiation zones. As possible reason of this result, they assumed that deposition zones mainly occur
413 close to bottom of small valleys which can be described by the selected terrain attributes better than
414 the open slope where initiation zones tend to be located. In other words, also Vorpahl et al. (2012)
415 assume that more specific terrain conditions of landslide data provide more accurate predictive skill
416 of the models.

417 Although apparently similar, the six debris-avalanche and -flow landslide susceptibility maps
418 obtained from MARS and LR models have important differences, which are revealed by the
419 histograms of Fig. 9. Based on the latter, we can infer that, also across the entire study area, MARS
420 performed better than LR and both modeling techniques achieved better prediction skill when
421 employed to discriminate between stable and unstable pixels of the dataset SOURCE. We consider
422 MARS maps better than LR maps because of the following reasons. First, frequency of
423 susceptibility levels of MARS maps gradually decreases from lowest to highest ones, which is a
424 desirable quality for landslide susceptibility maps. Indeed, as landsliding is normally a rare-event
425 and thus ratio of event to non-event cells can be very low, maps with high frequency of susceptible
426 pixels can suffer from high false positive rates. Second, using $P < 0.5$ as threshold to identify non-
427 susceptible pixels, MARS maps are characterized by significantly higher specificity (or true
428 negative rate) on all analyzed landslide datasets. Third, MARS maps have higher percentage of
429 event pixels with $P > 0.75$. Similarly, it is reasonable to infer that dataset SOURCE yielded better
430 MARS and LR maps because, although roughly the same true positive rate (P threshold = 0.5) was
431 observed for all landslide datasets, the percentage of event pixels with $P > 0.75$ was definitely
432 higher. Furthermore, a higher percentage of stable pixels with $P < 0.5$ occur in MARS- and LR-
433 SOURCE maps.

434 The results of this experiment demonstrated that a stochastic approach can be used by
435 geomorphologists to achieve reliable prediction of landslides triggered by an extreme rainfall event
436 in a mountainous environment. The approach described in this paper could be useful particularly in
437 developing countries which have great difficulty in affording the high costs of structural measures
438 to reduce landslide risk. This method allows for identifying areas prone to landsliding and thus
439 could provide a valuable aid to a rational land use planning aimed at minimizing victims and
440 economic damage.

441

442 **6. Concluding remarks**

443 In this experiment, we employed a stochastic approach to predict the spatial distribution of the
444 debris-avalanches and -flows triggered by heavy rainfalls occurred in the area of Mocoa on April
445 1st, 2017. Multivariate Adaptive Regression Splines (MARS) and Logistic Regression (LR) were
446 exploited as modeling techniques. A set of nine terrain attributes, in addition to lithology, distance
447 from faults and presence/absence of soil creep processes were used as predictor variables.
448 Validation of the models was performed by using k -fold spatial cross-validation and by calculating
449 AUC values of 100 MARS and LR model repetitions applied to balanced samples of event and non-
450 event pixels. The calibration and validation samples were extracted from three different landslide
451 datasets which contain the highest point of each landslide crown-line (dataset LIP), the highest 10%
452 of cells of each landslide area (dataset SOURCE) or the entire landslide area (dataset MASS).

453

454 The following conclusions can be drawn for the Mocoa study area on the basis of the validation of
455 the models and the analysis of six susceptibility maps which were obtained by using MARS and LR
456 with the three landslide datasets.

- 457 1. The overall accuracy of our models can be considered acceptable ($AUC > 0.7$);
- 458 2. MARS models and maps exhibited a better ability to discriminate between event and non-event
459 pixels;

460 3. Both MARS and LR demonstrated better accuracy when employed to predict landslide source
461 areas;

462 4. The relationships between predictors and debris-avalanche and -flow landslides are more
463 accurately reproduced by piecewise linear regressions rather than individual linear functions
464 holding over the entire predictor range.

465 The approach employed in this experiment is relatively simple, rapid and can be reproduced by
466 using free software and data usually available. It can be used to prepare debris-avalanche and -flow
467 landslide susceptibility maps which can help policy makers and land managers of Colombia to
468 establish preventive measures and mitigate risks.

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477

478

479 **References**

480 Aleotti, P., Chowdhury, R., 1999. Landslide hazard assessment: summary review and new perspectives. Bull.
481 Eng. Geol. Environ. 58, 21–44. doi:10.1007/s100640050066

482 Aronica, G.T., Brigandí, G., Morey, N., 2012. Flash floods and debris flow in the city area of Messina,
483 north-east part of Sicily, Italy in October 2009: The case of the Giampileri catchment. Nat. Hazards
484 Earth Syst. Sci. 12, 1295–1309. doi:10.5194/nhess-12-1295-2012

- 485 Atkinson, P.M., Massari, R., 2011. Autologistic modelling of susceptibility to landsliding in the Central
486 Apennines, Italy. *Geomorphology* 130, 55–64. doi:10.1016/j.geomorph.2011.02.001
- 487 Auslander, M., Nevo, E., Inbar, M., 2003. The effects of slope orientation on plant growth, developmental
488 instability and susceptibility to herbivores. *J. Arid Environ.* 55, 405–416. doi:10.1016/S0140-
489 1963(02)00281-1
- 490 Avelar, A.S., Netto, A.L.C., Lacerda, W.A., Becker, L.B., Mendonça, M.B., 2013. Mechanisms of the
491 Recent Catastrophic Landslides in the Mountainous Range of Rio de Janeiro, Brazil BT - Landslide
492 Science and Practice: Volume 4: Global Environmental Change, in: Margottini, C., Canuti, P., Sassa, K.
493 (Eds.), Springer Berlin Heidelberg, Berlin, Heidelberg, pp. 265–270. doi:10.1007/978-3-642-31337-
494 0_34
- 495 Beguería, S., 2006. Validation and Evaluation of Predictive Models in Hazard Assessment and Risk
496 Management. *Nat. Hazards* 37, 315–329. doi:10.1007/s11069-005-5182-6
- 497 Beven, K.J., Kirkby, M.J., 1979. A physically based variable contributing area model of basin hydrology.
498 *Hydrol. Sci. Bull.* 24, 43–69. doi:http://dx.doi.org/10.1080/02626667909491834
- 499 Brabb, E.E., 1984. Innovative approaches to landslide hazard and risk mapping, in: Proceedings 4th
500 International Symposium on Landslides Vol. 1. Toronto, Canada, pp. 307–324.
- 501 Brenning, A., 2005. Spatial prediction models for landslide hazards: review, comparison and evaluation. *Nat.*
502 *Hazards Earth Syst. Sci.* 5, 853–862. doi:10.5194/nhess-5-853-2005
- 503 Brenning, A., Trombotto, D., 2006. Logistic regression modeling of rock glacier and glacier distribution:
504 Topographic and climatic controls in the semi-arid Andes. *Geomorphology*.
505 doi:10.1016/j.geomorph.2006.04.003
- 506 Brenning, A., 2012. Improved spatial analysis and prediction of landslide susceptibility: Practical
507 recommendations. *Landslides Eng. Slopes Prot. Soc. through Improv. Understanding, Proc. 11th Int.*
508 *2nd North Am. Symp. Landslides Eng. Slopes, Banff, Canada* 789–794.
- 509 Bui, D.T., Tuan, T.A., Klempe, H., Pradhan, B., Revhaug, I., 2015. Spatial prediction models for shallow
510 landslide hazards : a comparative assessment of the efficacy of support vector machines , artificial

511 neural networks , kernel logistic regression , and logistic model tree. *Landslides*. doi:10.1007/s10346-
512 015-0557-6

513 Cama, M., Lombardo, L., Conoscenti, C., Rotigliano, E., 2017. Improving transferability strategies for debris
514 flow susceptibility assessment: Application to the Saponara and Itala catchments (Messina, Italy).
515 *Geomorphology* 288, 52–65. doi:10.1016/j.geomorph.2017.03.025

516 Cama, M., Conoscenti, C., Lombardo, L., Rotigliano, E., 2016. Exploring relationships between grid cell
517 size and accuracy for debris-flow susceptibility models: a test in the Giampileri catchment (Sicily,
518 Italy). *Environ. Earth Sci.* 75. doi:10.1007/s12665-015-5047-6

519 Carrara, A., Cardinali, M., Guzzetti, F., Reichenbach, P., 1995. GIS technology in mapping landslide hazard,
520 in: Carrara, A., Guzzetti, F. (Eds.), *Geographical Information Systems in Assessing Natural Hazards*.
521 Kluwer, Dordrecht, pp. 135–175.

522 Chevalier, G.G., Medina, V., Hürlimann, M., Bateman, A., 2013. Debris-flow susceptibility analysis using
523 fluvio-morphological parameters and data mining : *Nat. Hazards* 67, 213–238. doi:10.1007/s11069-013-
524 0568-3

525 Clerici, A., Perego, S., Tellini, C., Vescovi, P., 2006. A GIS-based automated procedure for landslide
526 susceptibility mapping by the Conditional Analysis method: the Baganza valley case study (Italian
527 Northern Apennines). *Environ. Geol.* 50, 941–961. doi:10.1007/s00254-006-0264-7

528 Conoscenti, C., Agnesi, V., Cama, M., Caraballo-Arias, N.A., Rotigliano, E., 2018. Assessment of Gully
529 Erosion Susceptibility Using Multivariate Adaptive Regression Splines and Accounting for Terrain
530 Connectivity. *L. Degrad. Dev.* 29, 724–736. doi:10.1002/ldr.2772

531 Conoscenti, C., Ciaccio, M., Caraballo-Arias, N.A., Gómez-Gutiérrez, Á., Rotigliano, E., Agnesi, V., 2015.
532 Assessment of susceptibility to earth-flow landslide using logistic regression and multivariate adaptive
533 regression splines: A case of the Belice River basin (western Sicily, Italy). *Geomorphology* 242, 49–64.
534 doi:10.1016/j.geomorph.2014.09.020

535 Conoscenti, C., Rotigliano, E., Cama, M., Caraballo-Arias, N.A., Lombardo, L., Agnesi, V., 2016. Exploring
536 the effect of absence selection on landslide susceptibility models: A case study in Sicily, Italy.
537 *Geomorphology* 261, 222–235. doi:10.1016/j.geomorph.2016.03.006

- 538 Conrad, O., Bechtel, B., Bock, M., Dietrich, H., Fischer, E., Gerlitz, L., Wehberg, J., Wichmann, V., Böhner,
539 J., 2015. System for Automated Geoscientific Analyses (SAGA) v. 2.1.4. *Geosci. Model Dev.* 8, 1991–
540 2007. doi:10.5194/gmd-8-1991-2015
- 541 Costanzo, D., Cappadonia, C., Conoscenti, C., Rotigliano, E., 2012. Exporting a Google Earth™ aided earth-
542 flow susceptibility model: a test in central Sicily. *Nat. Hazards* 61, 103–114. doi:10.1007/s11069-011-
543 9870-0
- 544 Cruden, D.M., Varnes, D.J., 1996. Landslide types and processes, in: Turner, A.K., Schuster, R.L. (Eds.),
545 Landslides: Investigation and Mitigation. Transportation Research Board Special Report, Vol. 247.
546 National Academy Press, Washington D.C., pp. 36–75.
- 547 Felicísimo, Á.M., Cuartero, A., Remondo, J., Quirós, E., 2013. Mapping landslide susceptibility with logistic
548 regression, multiple adaptive regression splines, classification and regression trees, and maximum
549 entropy methods: a comparative study. *Landslides* 10, 175–189. doi:10.1007/s10346-012-0320-1
- 550 Friedman, J.H., 1991. Multivariate adaptive regression splines. *Ann. Stat.* 19, 1–141.
- 551 Garosi, Y., Sheklabadi, M., Porghasemi, H.R., Besalatpour, A.A., Conoscenti, C., Van Oost, K., 2018.
552 Comparison of differences in resolution and sources of controlling factors for gully erosion
553 susceptibility mapping. *Geoderma* 330, 65–78. doi:10.1016/j.geoderma.2018.05.027
- 554 Goetz, J.N., Brenning, A., Petschko, H., Leopold, P., 2015. Evaluating machine learning and statistical
555 prediction techniques for landslide susceptibility modeling. *Comput. Geosci.* 81, 1–11.
556 doi:10.1016/j.cageo.2015.04.007
- 557 Gómez-Gutiérrez, Á., Conoscenti, C., Angileri, S.E., Rotigliano, E., Schnabel, S., 2015. Using topographical
558 attributes to evaluate gully erosion proneness (susceptibility) in two mediterranean basins: advantages
559 and limitations. *Nat. Hazards* 79, 291–314. doi:10.1007/s11069-015-1703-0
- 560 Gorum, T., Gonencgil, B., Gokceoglu, C., Nefeslioglu, H.A., 2008. Implementation of reconstructed
561 geomorphologic units in landslide susceptibility mapping: the Melen Gorge (NW Turkey). *Nat. Hazards*
562 46, 323–351. doi:10.1007/s11069-007-9190-6
- 563 Guisan, A., Weiss, S.B., Weiss, A.D., 1999. GLM versus CCA spatial modeling of plant species distribution.
564 *Plant Ecol.* 143, 107–122.

- 565 Guzzetti, F., Carrara, A., Cardinali, M., Reichenbach, P., 1999. Landslide hazard evaluation: a review of
566 current techniques and their application in a multi-scale study, Central Italy. *Geomorphology* 31, 181–
567 216. doi:10.1016/S0169-555X(99)00078-1
- 568 Heckmann, T., Gegg, K., Gegg, a., Becht, M., 2014. Sample size matters: investigating the effect of sample
569 size on a logistic regression susceptibility model for debris flows. *Nat. Hazards Earth Syst. Sci.* 14,
570 259–278. doi:10.5194/nhess-14-259-2014
- 571 Hosmer, D.W., Lemeshow, S., 2000. Applied logistic regression, Wiley Series in Probability and Statistics,
572 Wiley series in probability and statistics: Texts and references section. Wiley.
573 doi:10.1198/tech.2002.s650
- 574 Hungr, O., Leroueil, S., Picarelli, L., 2014. The Varnes classification of landslide types, an update.
575 *Landslides* 11, 167–194. doi:10.1007/s10346-013-0436-y
- 576 Jebur, M.N., Pradhan, B., Tehrany, M.S., 2014. Optimization of landslide conditioning factors using very
577 high-resolution airborne laser scanning (LiDAR) data at catchment scale. *Remote Sens. Environ.* 152, 150–
578 165. doi:10.1016/j.rse.2014.05.013
- 579 Köthe, R., Lehmeier, F., 1996. SARA - System zur Automatischen Relief-Analyse, Benutzerhandbuch.
580 Department of Geography, University of Göttingen, Göttingen.
- 581 Larsen, M.C., Wiczorek, G.F., 2006. Geomorphic effects of large debris flows and flash floods, northern
582 Venezuela, 1999. *Zeitschrift für Geomorphol.* 145, 147–175.
- 583 Larsen, M.C., 2008. Rainfall-triggered landslides, anthropogenic hazards, and mitigation strategies. *Adv.*
584 *Geosci.* 14, 147–153. doi:10.5194/adgeo-14-147-2008
- 585 Lóczy, D., Pirkhoffer, E., Gyenizse, P., 2012. Geomorphometric floodplain classification in a hill region of
586 Hungary. *Geomorphology* 147–148, 61–72. doi:10.1016/j.geomorph.2011.06.040
- 587 Lombardo, L., Cama, M., Conoscenti, C., Märker, M., Rotigliano, E., 2015. Binary logistic regression versus
588 stochastic gradient boosted decision trees in assessing landslide susceptibility for multiple-occurring
589 landslide events: application to the 2009 storm event in Messina (Sicily, southern Italy). *Nat. Hazards*
590 79, 1621–1648. doi:10.1007/s11069-015-1915-3

591 Lombardo, L., Cama, M., Maerker, M., Rotigliano, E., 2014. A test of transferability for landslides
592 susceptibility models under extreme climatic events: application to the Messina 2009 disaster. *Nat.*
593 *Hazards* 1951–1989. doi:10.1007/s11069-014-1285-2

594 Mathew, J., Jha, V.K., Rawat, G.S., 2009. Landslide susceptibility zonation mapping and its validation in
595 part of Garhwal Lesser Himalaya, India, using binary logistic regression analysis and receiver operating
596 characteristic curve method. *Landslides* 6, 17–26. doi:10.1007/s10346-008-0138-z

597 Milborrow, S., 2015. Notes on the earth package [WWW Document]. URL [http://www.milbo.org/doc/earth-](http://www.milbo.org/doc/earth-notes.pdf)
598 [notes.pdf](http://www.milbo.org/doc/earth-notes.pdf) (accessed 7.15.15).

599 Naimi, B., 2015. Uncertainty analysis for species distribution models. R Software Package.

600 Nefeslioglu, H.A., Gokceoglu, C., Sonmez, H., 2008. An assessment on the use of logistic regression and
601 artificial neural networks with different sampling strategies for the preparation of landslide
602 susceptibility maps. *Eng. Geol.* 97, 171–191. doi:10.1016/j.enggeo.2008.01.004

603 Nicu, I.C., Asăndulesei, A., 2018. GIS-based evaluation of diagnostic areas in landslide susceptibility
604 analysis of Bahluieț River Basin (Moldavian Plateau, NE Romania). Are Neolithic sites in danger?
605 *Geomorphology* 314, 27–41. doi:10.1016/j.geomorph.2018.04.010

606 Ohlmacher, G.C., 2007. Plan curvature and landslide probability in regions dominated by earth flows and
607 earth slides. *Eng. Geol.* 91, 117–134. doi:10.1016/j.enggeo.2007.01.005

608 Pourghasemi, H.R., Rossi, M., 2016. Landslide susceptibility modeling in a landslide prone area in
609 Mazandarn Province, north of Iran: a comparison between GLM, GAM, MARS, and M-AHP methods.
610 *Theor. Appl. Climatol.* doi:10.1007/s00704-016-1919-2

611 R Core Team, 2017. R: A Language and Environment for Statistical Computing.

612 Regmi, N.R., Giardino, J.R., McDonald, E. V., Vitek, J.D., 2014. A comparison of logistic regression-based
613 models of susceptibility to landslides in western Colorado, USA. *Landslides* 11, 247–262.
614 doi:10.1007/s10346-012-0380-2

615 Reichenbach, P., Rossi, M., Malamud, B.D., Mihir, M., Guzzetti, F., 2018. A review of statistically-based
616 landslide susceptibility models. *Earth-Science Rev.* 180, 60–91. doi:10.1016/j.earscirev.2018.03.001

- 617 Rotigliano, E., Agnesi, V., Cappadonia, C., Conoscenti, C., 2011. The role of the diagnostic areas in the
618 assessment of landslide susceptibility models: a test in the sicilian chain. *Nat. Hazards* 58, 981–999.
619 doi:10.1007/s11069-010-9708-1
- 620 Rotigliano, E., Martinello, C., Agnesi, V., Conoscenti, C., 2018. Evaluation of debris flow susceptibility in
621 El Salvador (CA): a comparison between Multivariate Adaptive Regression Splines (MARS) and
622 Binary Logistic Regression (BLR). *Hungarian Geogr. Bull.* 67, 361–373.
623 doi:10.15201/hungeobull.67.4.5
- 624 Ruß, G., Brenning, A., 2010. *Computational Intelligence for Knowledge-Based Systems Design, Lecture*
625 *Notes in Computer Science (including subseries Lecture Notes in Artificial Intelligence and Lecture*
626 *Notes in Bioinformatics), Lecture Notes in Computer Science.* Springer Berlin Heidelberg, Berlin,
627 Heidelberg. doi:10.1007/978-3-642-14049-5
- 628 Süzen, M.L., Doyuran, V., 2004. Data driven bivariate landslide susceptibility assessment using
629 geographical information systems: a method and application to Asarsuyu catchment, Turkey. *Eng. Geol.*
630 71, 303–321. doi:10.1016/S0013-7952(03)00143-1
- 631 Svoray, T., Michailov, E., Cohen, A., Rokah, L., Sturm, A., 2012. Predicting gully initiation: comparing data
632 mining techniques, analytical hierarchy processes and the topographic threshold. *Earth Surf. Process.*
633 *Landforms* 37, 607–619. doi:10.1002/esp.2273
- 634 Trigila, A., Iadanza, C., Esposito, C., Scarascia-Mugnozza, G., 2015. Comparison of Logistic Regression and
635 Random Forests techniques for shallow landslide susceptibility assessment in Giampileri (NE Sicily,
636 Italy). *Geomorphology*. <https://doi.org/10.1016/j.geomorph.2015.06.001>
- 637 Van Den Eeckhaut, M., Hervás, J., Jaedicke, C., Malet, J.P., Montanarella, L., Nadim, F., 2012. Statistical
638 modelling of Europe-wide landslide susceptibility using limited landslide inventory data. *Landslides* 9,
639 357–369. doi:10.1007/s10346-011-0299-z
- 640 Van Den Eeckhaut, M., Vanwallegem, T., Poesen, J., Govers, G., Verstraeten, G., Vandekerckhove, L.,
641 2006. Prediction of landslide susceptibility using rare events logistic regression: A case-study in the
642 Flemish Ardennes (Belgium). *Geomorphology* 76, 392–410. doi:10.1016/j.geomorph.2005.12.003

- 643 Vorpahl, P., Elsenbeer, H., Märker, M., Schröder, B., 2012. How can statistical models help to determine
644 driving factors of landslides? *Ecol. Modell.* 239, 27–39. doi:10.1016/j.ecolmodel.2011.12.007
- 645 Wilson, J.P., Gallant, J.C., 2000. *Terrain Analysis: Principles and Applications*. Wiley & Sons, Inc., Canada.
- 646 Zevenbergen, L.W., Thorne, C.R., 1987. Quantitative analysis of land surface topography. *Earth Surf.*
647 *Process. Landforms* 12, 47–56.

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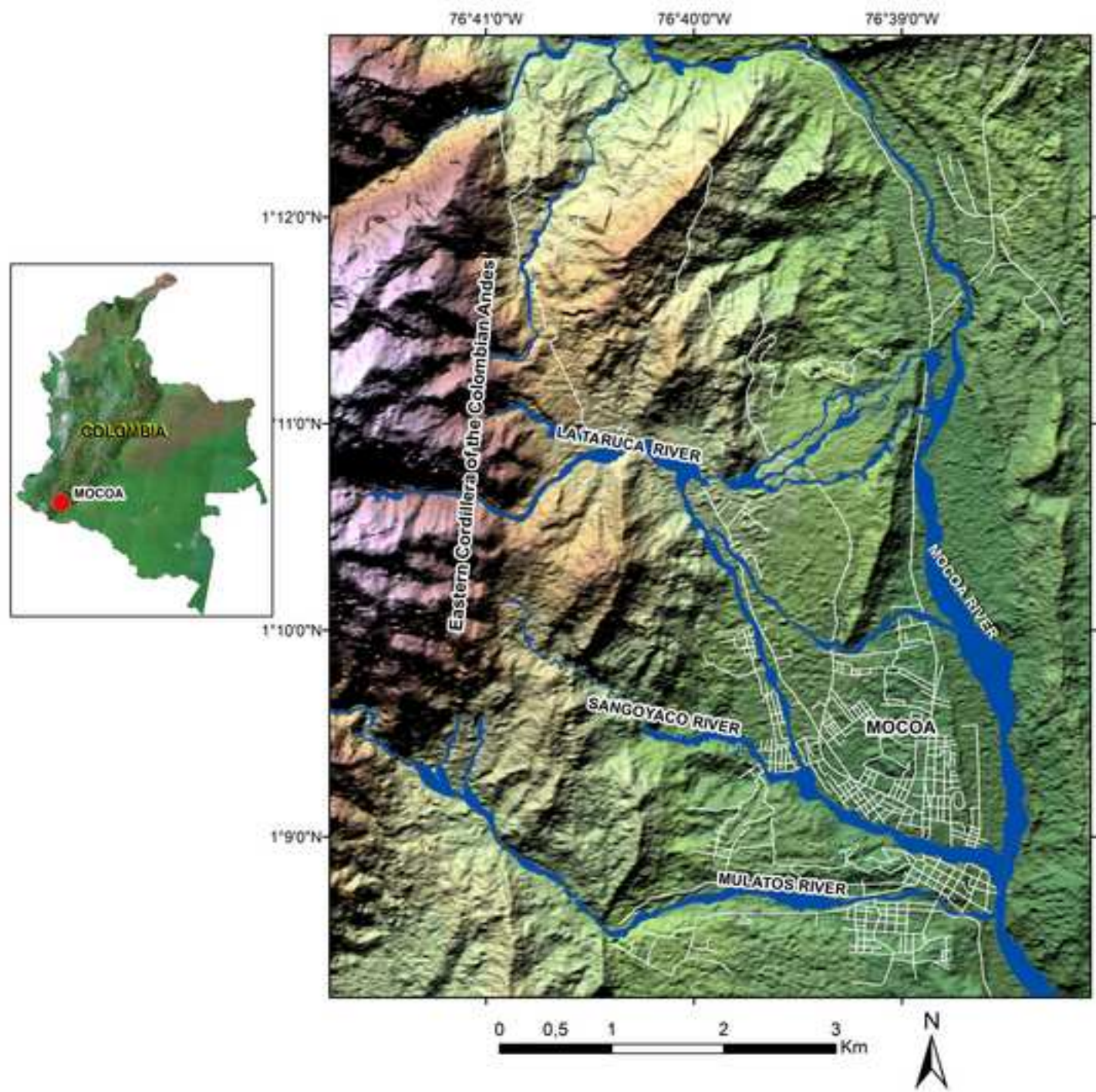


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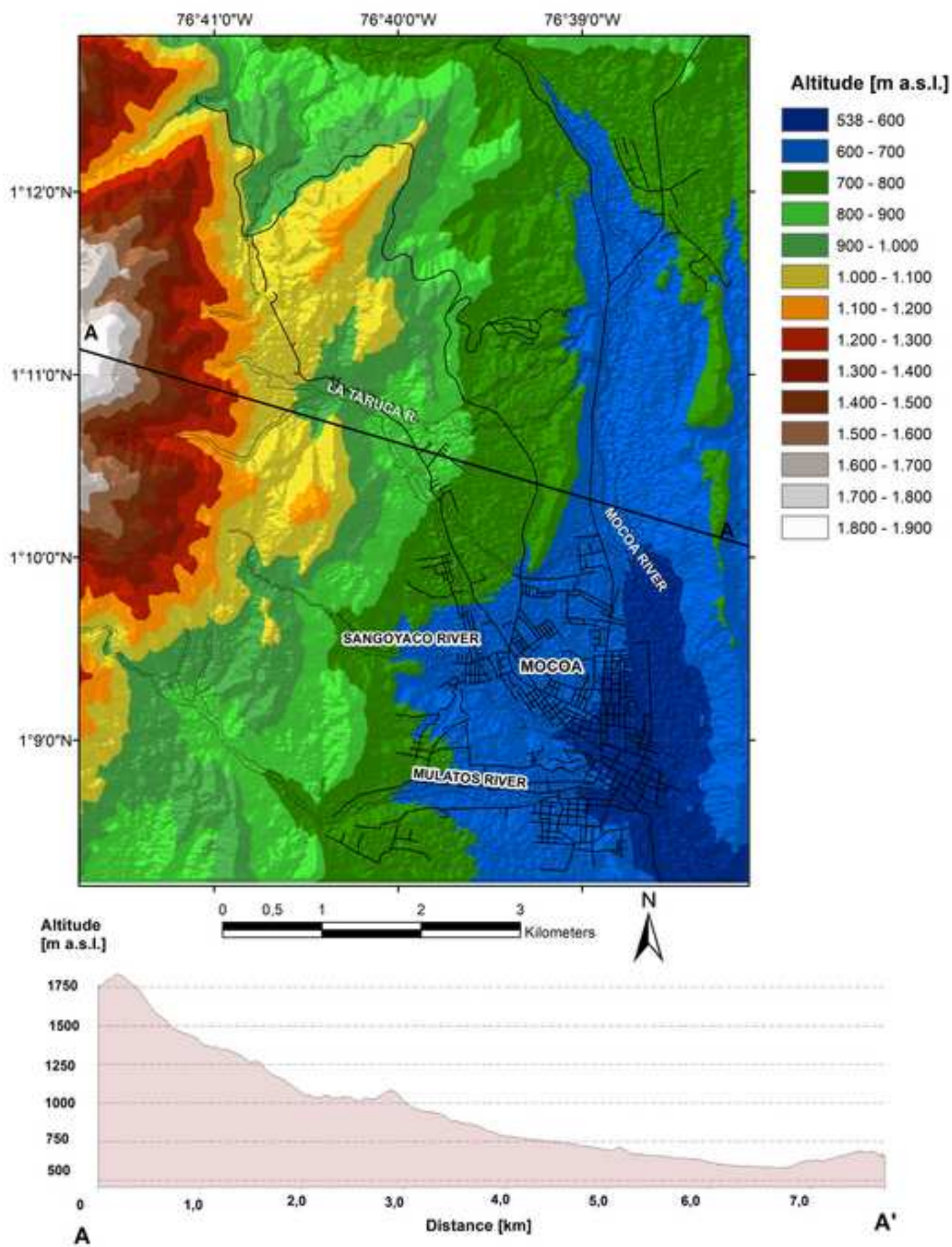


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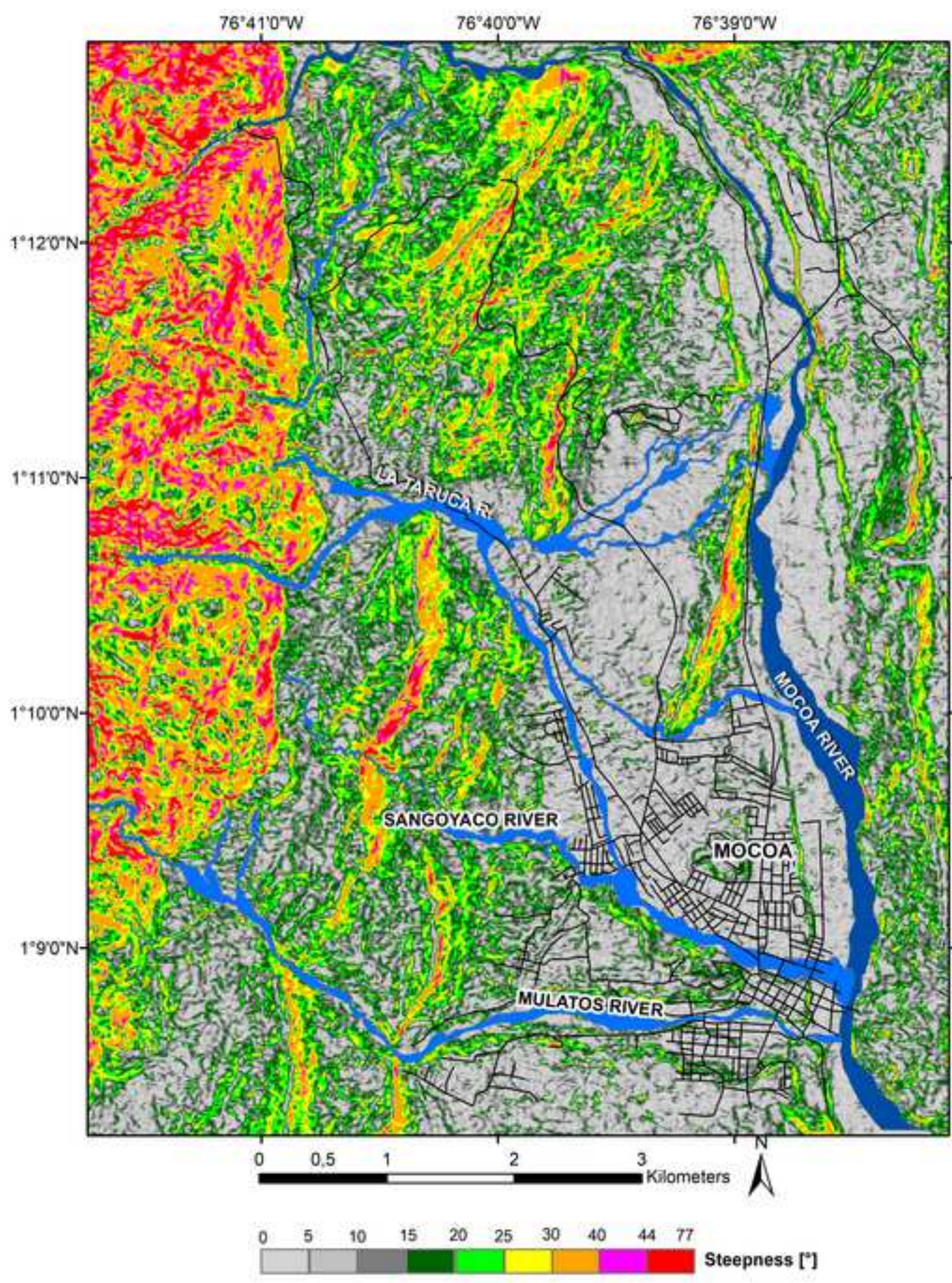


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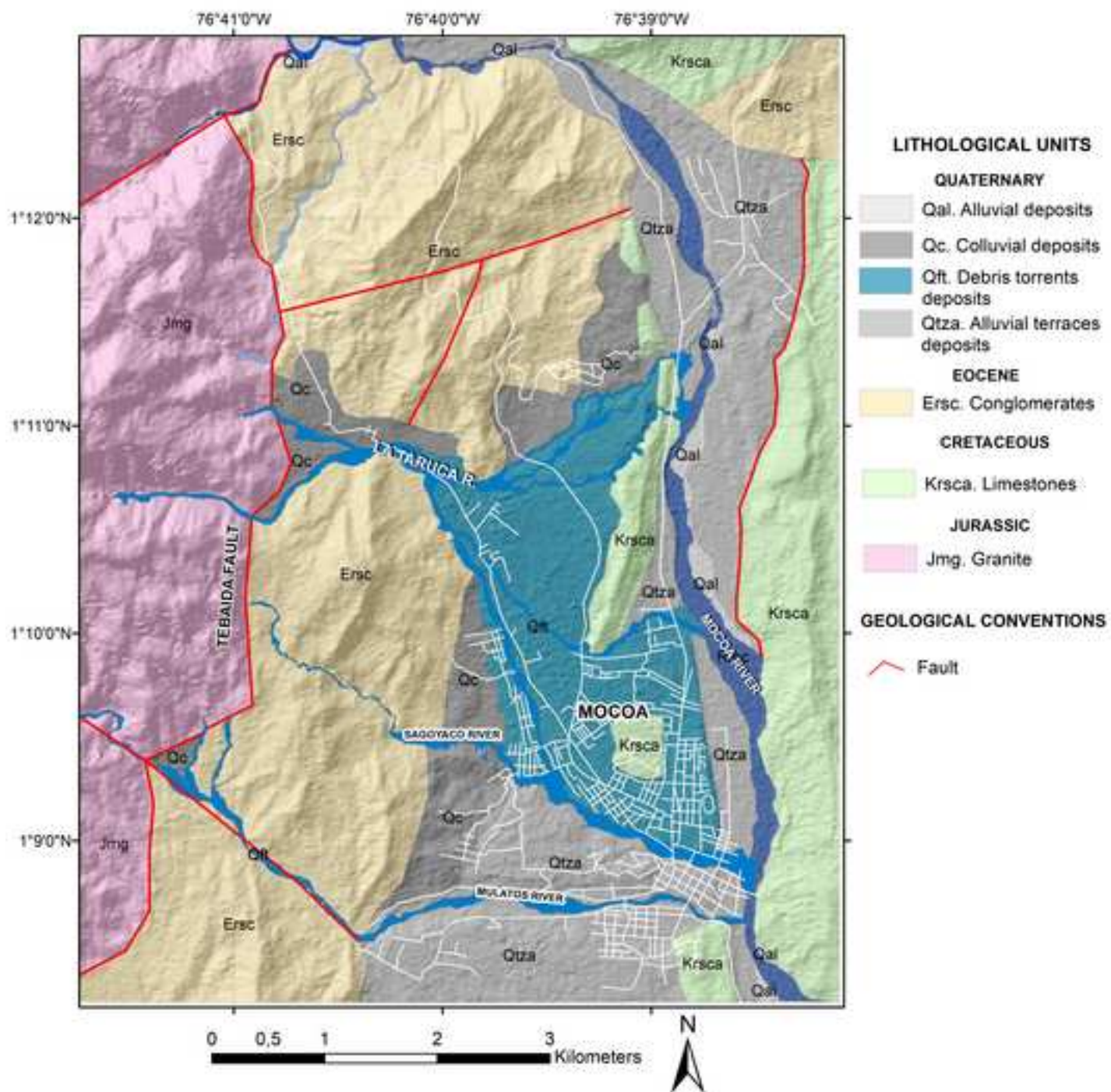


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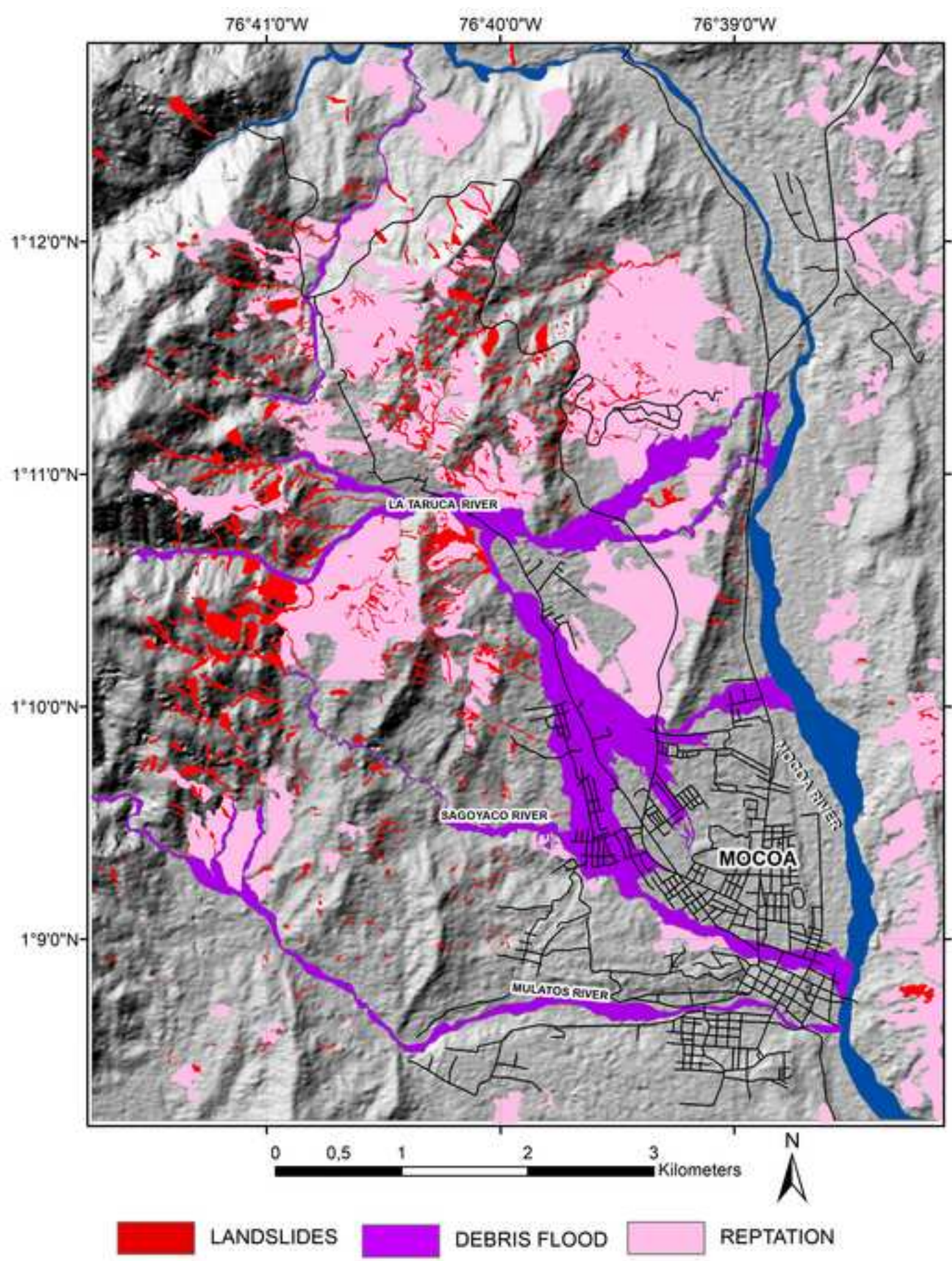


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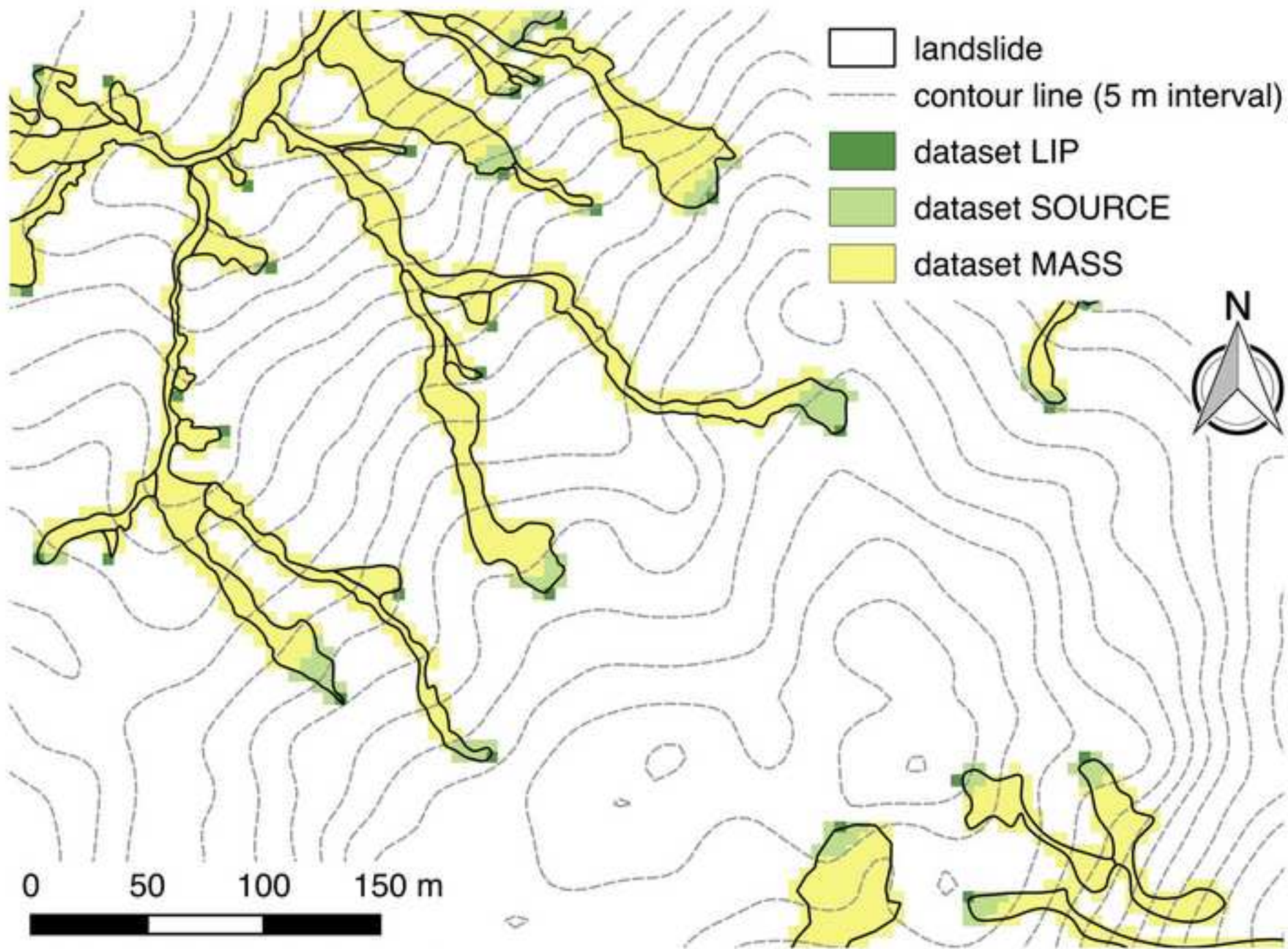


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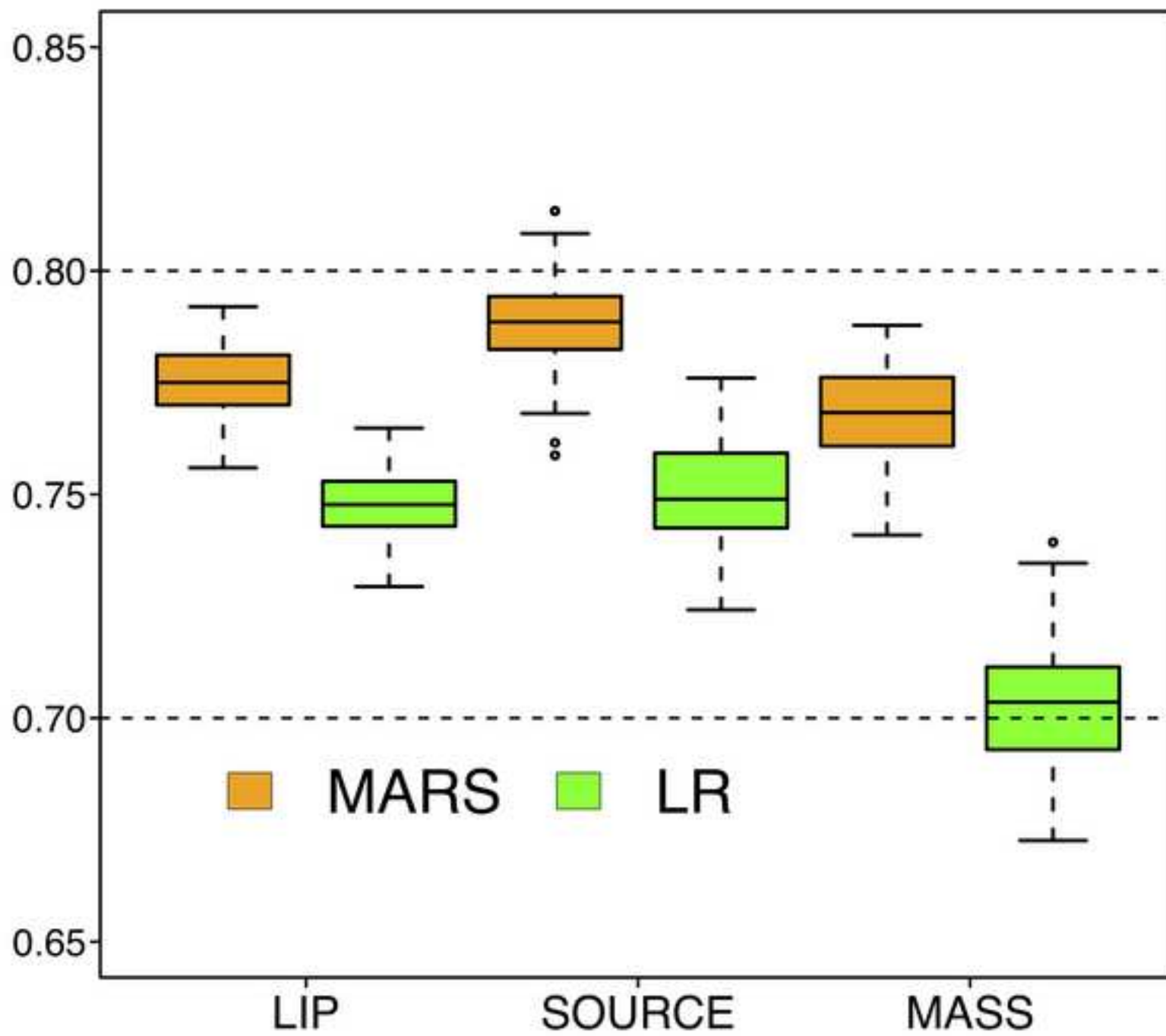


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